



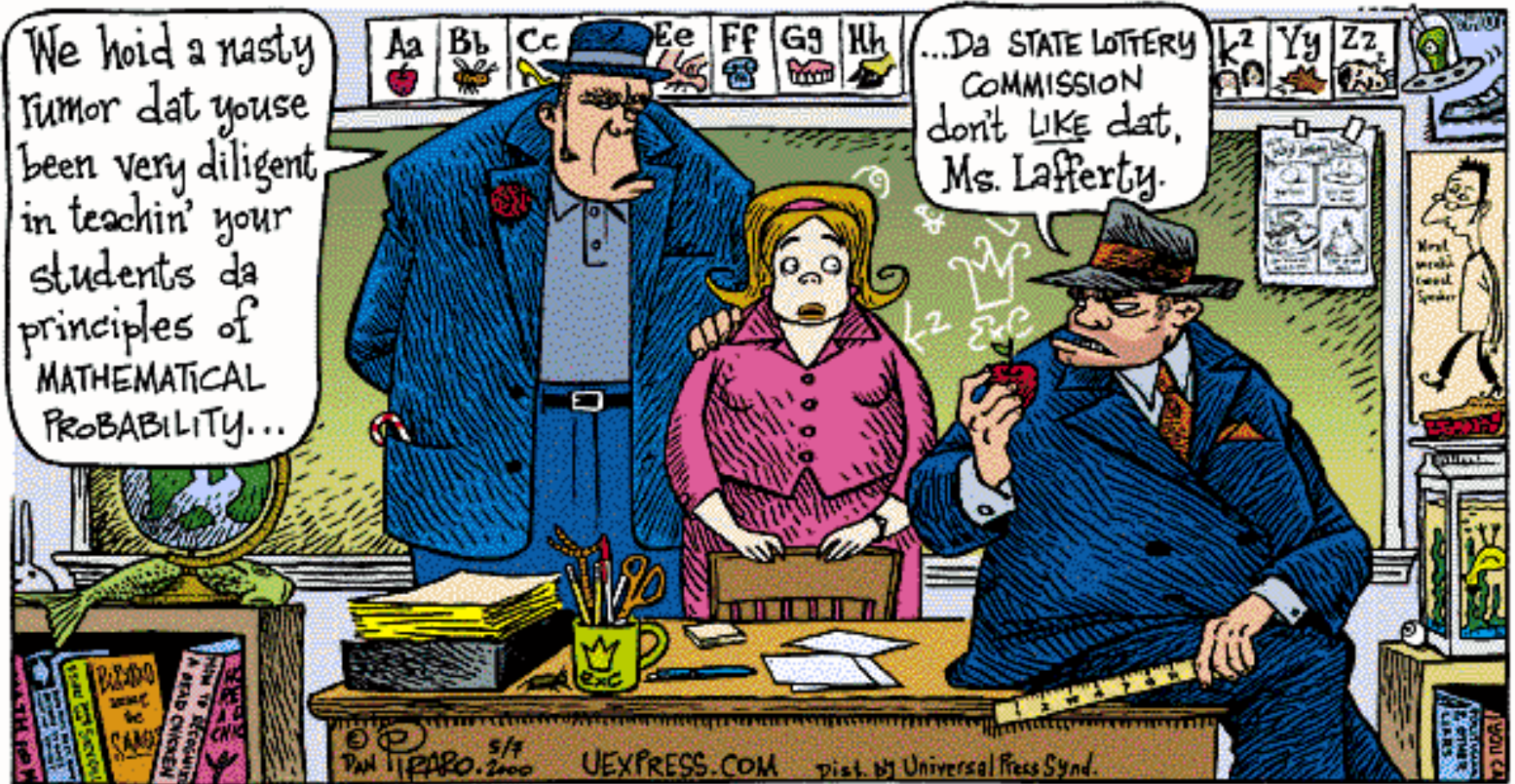
Algorithmic Abstractions of (Probabilistic) Reasoning: Bayesian Networks





Probabilistic Knowledge Representation

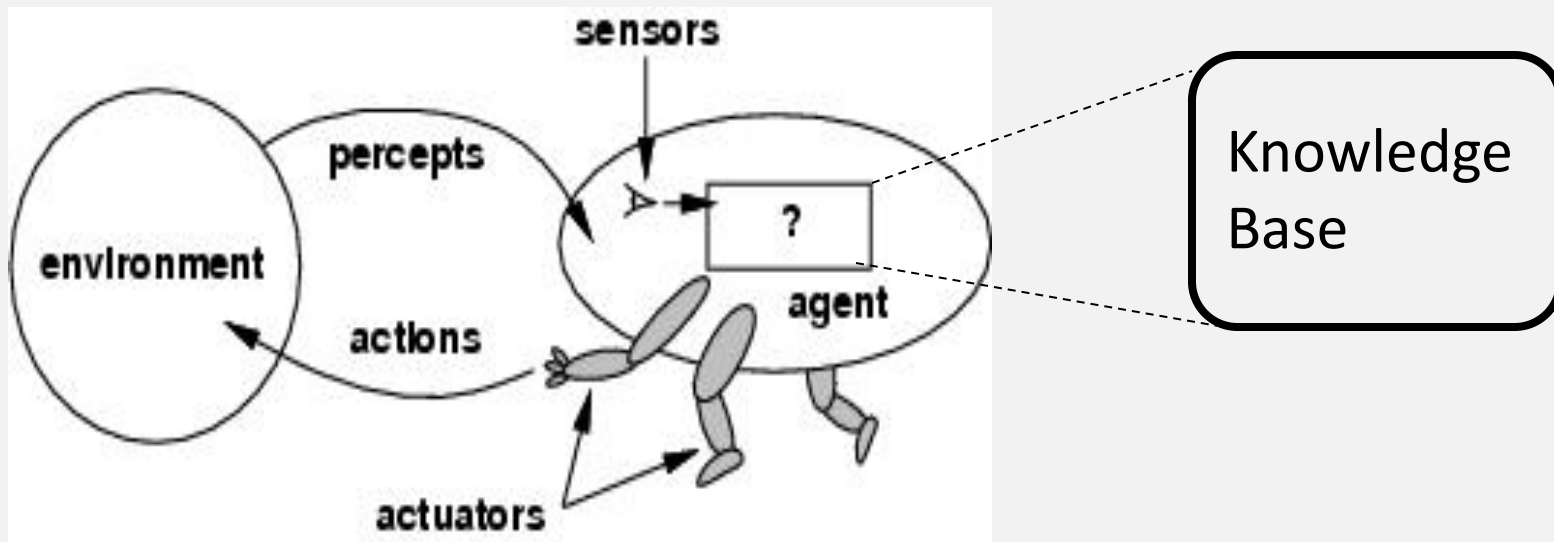
- Basic probability theory
- Syntax and Semantics
- Random variables
- Distributions over random variables
- Independence and conditional independence
- Bayesian Network Representation
- Inference Using Bayesian Networks





Agents That Represent and Reason Under Uncertainty

- Intelligent behavior requires knowledge about the world
- Often, we are uncertain about the state of the world





Representing and Reasoning under Uncertainty

- Probability Theory provides a framework for representing and reasoning under uncertainty
 - Represent **beliefs** about **the world** as **sentences** (much like in **propositional logic**)
 - Associate **probabilities** with sentences
 - **Reason** by manipulating **sentences** according to **sound** rules of **probabilistic inference**
 - Results of inference are probabilities associated with conclusions that are justified by beliefs and data (observations)
- Allows agents to **substitute thinking for acting** in the world

Representing and Reasoning under Uncertainty

- Beliefs:
 - If Oksana studies, there is an 60% chance that she will pass the test; and a 40 percent chance that she will not.
 - If she does not study, there is 20% percent chance that she will pass the test and 80% chance that she will not.
- Observation: Oksana did not study.
- Example Inference task:
 - What is the chance that Oksana will pass the test?
 - What is the chance that she will fail?
- Probability theory generalizes propositional logic
 - Probability theory associates probabilities that lie in the interval $[0,1]$ as opposed to 0 or 1 (exclusively)



Probability Theory as a Knowledge Representation

- **Ontological commitments** (what do we want to talk about?)
 - Propositions that represent the agent's beliefs about the world
- **Epistemological Commitments** (what can we believe?)
 - What is the *probability* that a given proposition true (given the beliefs and observations)?
- **Syntax**
 - Much like propositional logic
- **Semantics**
 - Relative frequency interpretation
 - Bayesian interpretation
- **Proof Theory**
 - Based on laws of probability



Sources of uncertainty

Uncertainty modeled by Probabilistic assertions may

- In a deterministic world be due to
 - **Laziness**: failure to enumerate exceptions, qualifications, etc. that may be too numerous to state explicitly
 - Sensory limitations
 - **Ignorance**: lack of relevant facts etc.
- In a stochastic world be due to
 - Inherent uncertainty (as in quantum physics)

The framework is agnostic about the source of uncertainty



The world according to Agent Bob

- An **atomic event** or **world state** is a **complete specification** of the state of the agent's world.
- **Event set** is a set of mutually exclusive and exhaustive possible world states (relative to an agent's representational commitments and sensing abilities)
- **From the point of view of an agent Bob** who can sense only **3 colors** and **2 shapes**, the world can be in only one of **6 states**
- Atomic events (world states) are
 - **mutually exclusive**
 - **exhaustive**





Semantics: Probability as a subjective measure of belief

- Suppose there are 3 agents – Oksana, Cornelia, Jun, in a world where a fair dice has been tossed.
- Oksana observes that the outcome is a “6” and whispers to Cornelia that the outcome is “even” but
- Jun knows nothing about the outcome.

Set of possible mutually exclusive and exhaustive world states

$$= \{1, 2, 3, 4, 5, 6\}$$

Set of possible states of the world based on what Cornelia

$$\text{knows} = \{2, 4, 6\}$$





Probability as a subjective measure of belief

Probability is a **measure over all of the world states that are possible**, or simply, possible worlds, **given what an agent knows**

$$\text{Possibleworlds}_{\text{Oksana}} = \{6\}, \text{Possibleworlds}_{\text{Cornelia}} = \{2,4,6\}$$

$$\text{Possibleworlds}_{\text{Jun}} = \{1,2,3,4,5,6\}$$

$$\Pr_{\text{Oksana}}(\text{worldstate} = 6) = 1$$

$$\Pr_{\text{Cornelia}}(\text{worldstate} = 6) = \frac{1}{3}$$

$$\Pr_{\text{Jun}}(\text{worldstate} = 6) = \frac{1}{6}$$

Oksana, Cornelia, and Jun assign different beliefs to the same world state because of differences in their knowledge!





Random variables

- The “domain” of a random variable is the set of values it can take. The values are mutually exclusive and exhaustive.
- The domain of a Boolean random variable X is $\{\text{true}, \text{false}\}$ or $\{1, 0\}$
- **Discrete random variables** take values from a **countable** domain.
 - The domain of the random variable Color may be $\{\text{Red}, \text{Green}\}$.
 - If $E = \{(\text{Red}, \text{Square}), (\text{Green}, \text{Circle}), (\text{Red}, \text{Circle}), (\text{Green}, \text{Square})\}$, the proposition $(\text{Color} = \text{Red})$ is True in the world states $\{(\text{Red}, \text{Square}), (\text{Red}, \text{Circle})\}$.
 - Each state of a discrete random variable corresponds to a proposition e.g., $(\text{Color} = \text{Red})$



Syntax

- Basic element: **random variable**
 - Similar to propositional logic: possible worlds defined by assignment of values to random variables.
 - *Cavity* (do I have a cavity?)
 - *Weather* is one of $\langle \textit{sunny}, \textit{rainy}, \textit{cloudy}, \textit{snow} \rangle$
 - Domain values must be **exhaustive** and **mutually exclusive**
- Elementary proposition constructed by assignment of a value to a random variable
 - *Weather* = *sunny* (abbreviated as *sunny*), *Cavity* = *false* (abbreviated as $\neg \textit{cavity}$)
- Complex propositions formed from elementary propositions and standard logical connectives
 - *Weather* = *sunny* \vee $\neg \textit{cavity}$





Syntax and Semantics

- **Atomic event:** A **complete** specification of the state of the world about which the agent is uncertain
- Atomic events correspond to a possible worlds (much like in the case of propositional logic)

E.g., if the world consists of only two Boolean variables *Cavity* and *Toothache*, then there are 4 distinct atomic events or 4 possible worlds:

$$Cavity = false \wedge Toothache = false$$

$$Cavity = false \wedge Toothache = true$$

$$Cavity = true \wedge Toothache = false$$

$$Cavity = true \wedge Toothache = true$$

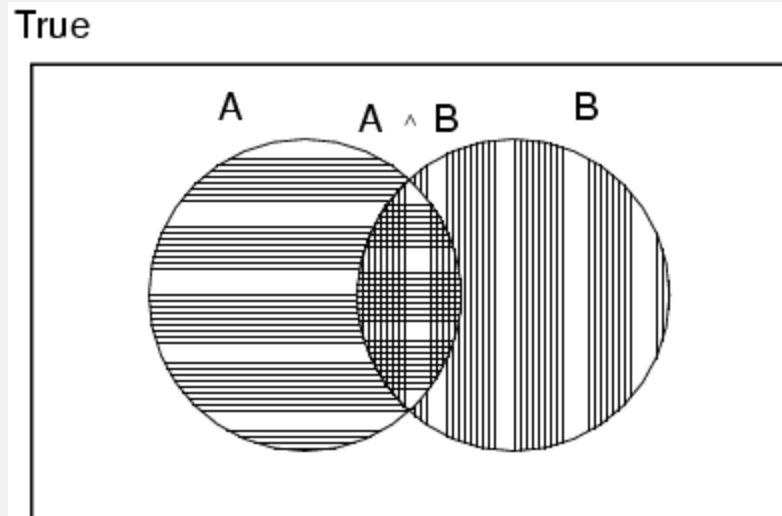
- Atomic events are mutually exclusive and exhaustive





Axioms of probability

- For any propositions A, B
 - $0 \leq P(A) \leq 1$
 - $P(\text{true}) = 1$ and $P(\text{false}) = 0$
 - $P(A \vee B) = P(A) + P(B) - P(A \wedge B)$





Prior probability

- **Prior** or **unconditional probabilities** of propositions
 - $P(\text{Cavity} = \text{true}) = 0.1$ and $P(\text{Weather} = \text{sunny}) = 0.72$
correspond to belief prior to arrival of any (new) evidence
- **Probability distribution** gives values for all possible assignments:
 - $\mathbf{P}(\text{Weather}) = \langle 0.72, 0.1, 0.08, 0.1 \rangle$
 - Note that the probabilities sum to 1
- **Joint probability distribution** for a set of random variables gives the probability of every atomic event on those random variables
 - $\mathbf{P}(\text{Cavity}, \text{Play}) =$ a 4×2 matrix of values

Joint probability distribution

- **Joint probability distribution** for a set of random variables gives the probability of every atomic event on those random variables

- $P(\text{Weather}, \text{Cavity}) =$ a 4×2 matrix of values:

<i>Weather =</i>	sunny	rainy	cloudy	snow
<i>Cavity = true</i>	0.144	0.02	0.016	0.02
<i>Cavity = false</i>	0.576	0.08	0.064	0.08

- **Every question about a domain can be answered by the joint distribution**
-



Inference using the joint distribution

	Toothache	\neg Toothache
Cavity	0.4	0.1
\neg Cavity	0.1	0.4

$$P(\text{cavity}) = P(\text{cavity}, \text{ache}) + P(\text{cavity}, \neg \text{ache})$$



Conditional probability

- **Conditional** or **posterior probabilities**

- $P(\text{Cavity} \mid \text{Toothache}) = 0.8$

- (note *Cavity* is shorthand for $\text{Cavity} = \text{True}$)

- Probability of *Cavity* **given** *Toothache*

- Notation for conditional distributions:

- $\mathbf{P}(\text{Cavity} \mid \text{Toothache}) = 2\text{-element vector of } 2\text{-element vectors}$

- $P(\text{Cavity} \mid \text{Toothache}, \text{Cavity}) = 1$

- New evidence may be irrelevant (Probability of Cavity given Toothache is independent of Weather)

- $P(\text{Cavity} \mid \text{Toothache}, \text{Sunny}) = P(\text{Cavity} \mid \text{Toothache}) = 0.8$





Conditional probability

- Definition of conditional probability:
$$P(a \mid b) = P(a \wedge b) / P(b) \text{ if } P(b) > 0$$
- **Product rule** gives an alternative formulation:
 - $P(a \wedge b) = P(a \mid b) P(b) = P(b \mid a) P(a)$

Example:

- Suppose I have two coins – one a normal fair coin, and the other a rigged coin (with heads on both sides). I pick a coin at *random*, *toss it*, and tell you that the outcome of the toss is a Head.
- What is the probability that I am looking at a fair coin?



Conditional probability

Example:

- Suppose I have two coins – one a normal fair coin, and the other a rigged coin (with heads on both sides). I pick a coin at *random*, *toss it*, and tell you that the outcome of the toss is a Head.
- What is the probability that I am looking at a fair coin?
- (F, H), (F,T),(R,H), (R,T)

$\frac{1}{4}$, $\frac{1}{4}$, $\frac{1}{2}$, 0

$$P(F | H) = P(F,H)/P(H)=(1/4)/(3/4) = 1/3$$





Conditional probability

- A general version holds for whole distributions, e.g.,

$$\mathbf{P}(\textit{Weather}, \textit{Cavity}) = \mathbf{P}(\textit{Weather} \mid \textit{Cavity}) \mathbf{P}(\textit{Cavity})$$

- View as a compact notation for a set of 4×2 equations, **not** matrix multiplication

- **Chain rule** is derived by successive application of product rule:

$$\begin{aligned} \mathbf{P}(X_1, \dots, X_n) &= \mathbf{P}(X_1, \dots, X_{n-1}) \mathbf{P}(X_n \mid X_1, \dots, X_{n-1}) \\ &= \mathbf{P}(X_1, \dots, X_{n-2}) \mathbf{P}(X_{n-1} \mid X_1, \dots, X_{n-2}) \mathbf{P}(X_n \mid X_1, \dots, X_{n-1}) \\ &= \dots \\ &= \pi_i \mathbf{P}(X_i \mid X_1, \dots, X_{i-1}) \text{ (i ranges from 1 to n)} \end{aligned}$$



Possible worlds semantics

- A possible world is an assignment of Truth values to every simple proposition about the world. Let Ω be a set of possible worlds. Let $\omega \in \Omega$ and let p, q be propositions (atomic sentences or syntactically well formed logical formulae). Then p is True in ω (written $\omega \models p$) where

$\omega \models p$ if ω assigns value *True* to p

$\omega \models p \wedge q$ if $\omega \models p$ and $\omega \models q$

$\omega \models p \vee q$ if $\omega \models p$ or $\omega \models q$ (or both)

$\omega \models \neg p$ if $\omega \not\models p$



Possible Worlds and Random Variables

- A possible world is an assignment of exactly one value to every random variable. Let Ω be a set of possible worlds. Let $\omega \in \Omega$ and let f be a (logical) formula. Then f is True in ω (written $\omega \models f$) where

$\omega \models X = v$ if ω assigns value v to X

$\omega \models f \wedge g$ if $\omega \models f$ and $\omega \models g$

$\omega \models f \vee g$ if $\omega \models f$ or $\omega \models g$ (or both)

$\omega \models \neg f$ if $\omega \not\models f$



Probability as a Measure over Possible worlds

- Associated with each possible world is a measure. When there are only a finite number of possible worlds, the measure of the world ω , denoted by $\mu(\omega)$ has the following properties:

$$\forall \omega \in \Omega, 0 \leq \mu(\omega)$$

$$\sum_{\omega \in \Omega} \mu(\omega) = 1$$

The probability of a formula or state of affairs described by a sentence f , written as $P(f)$, is the sum of the measures of the possible worlds in which f is True. That is,

$$P(f) = \sum_{\omega \models f} \mu(\omega)$$





Probability as a measure over possible worlds

- Suppose I have two coins – one a normal fair coin, and the other with 2 heads. I pick a coin at *random* and toss it. What is the probability that coin I picked is Fair given the outcome is a Head?

$$\Omega = \{(Fair, H), (Fair, T), (Rigged, H), (Rigged, T)\}$$

$$\mu = \left\{ \frac{1}{4}, \frac{1}{4}, \frac{1}{2}, 0 \right\}$$

$$\Pr(H) = \sum_{\omega|H} \mu(\omega) = \frac{1}{4} + \frac{1}{2} = \frac{3}{4}$$

$$P(Fair|H) = \frac{P(Fair, H)}{P(H)} = \frac{\frac{1}{4}}{\frac{3}{4}} = 1/3$$





Conditional probability as a Measure over Possible worlds not ruled out by evidence

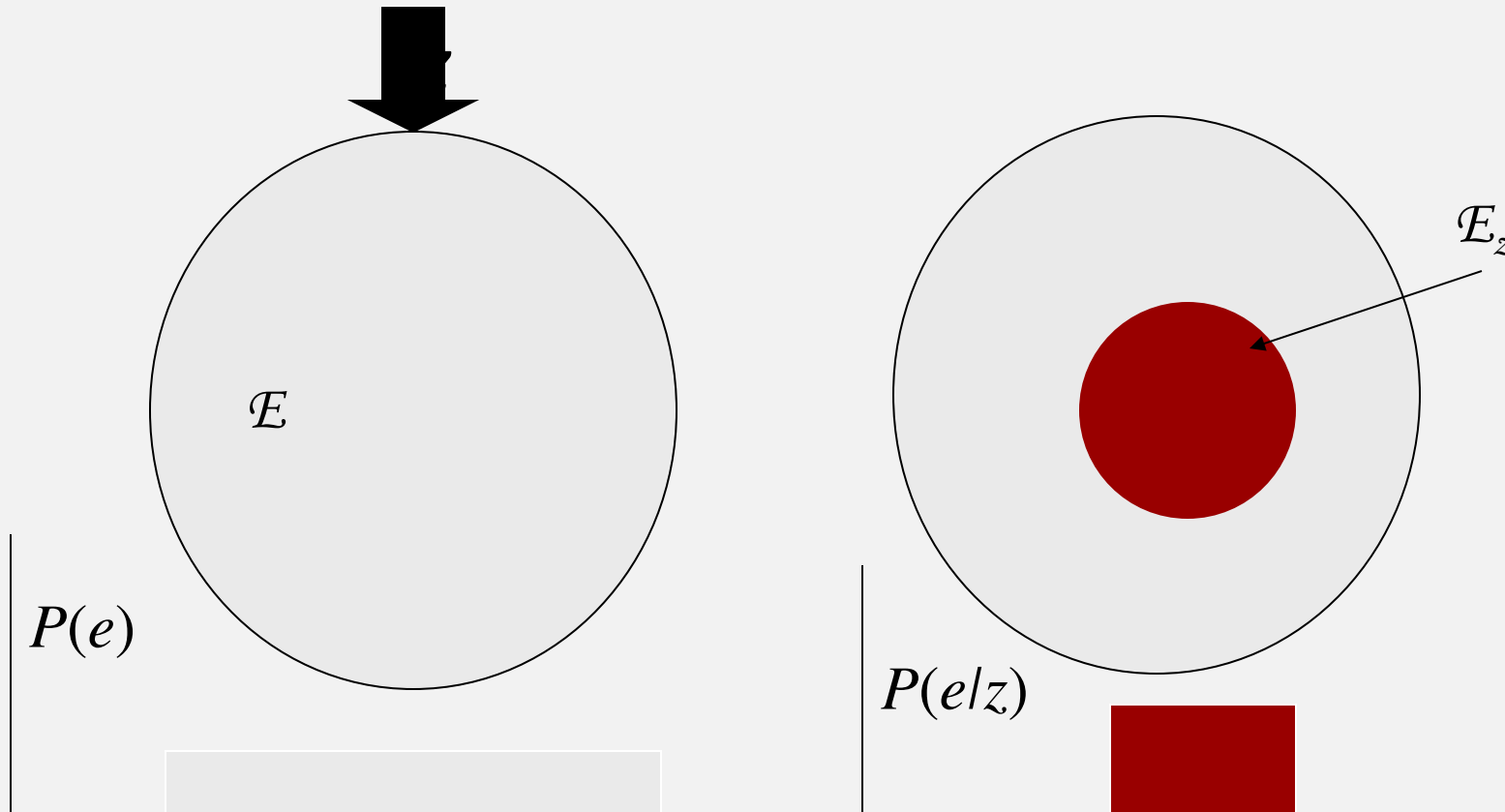
- A given piece of evidence e rules out all possible worlds that are incompatible with e or selects the possible worlds in which e is *True*. Evidence e induces a new measure μ_e .

$$\mu_e(\omega) = \begin{cases} \frac{1}{P(e)} \mu(\omega) & \text{if } \omega \models e \\ 0 & \text{if } \omega \not\models e \end{cases}$$

$$P(h|e) = \sum_{\omega \models h} \mu_e(\omega) = \frac{1}{P(e)} \sum_{\omega \models h \wedge e} \mu(\omega) = \frac{P(h \wedge e)}{P(e)}$$

Effect of Evidence on Possible worlds

Evidence z e.g., (color = red) rules out some assignments of values to some of the random variables





Evidence redistributes probability mass over possible worlds

- A given piece of evidence z rules out all possible worlds that are incompatible with z or selects the possible worlds in which z is *True*. Evidence z *induces* a distribution P_z

$$P_z(e) = \begin{cases} \frac{1}{P(z)} P(e) & \text{if } e \models z \\ 0 & \text{if } e \not\models z \end{cases}$$

$$P(h|z) = \sum_{e \models h} P_z(e) = \frac{1}{P(z)} \sum_{e \models h \wedge z} P(e) = \frac{P(h \wedge z)}{P(z)}$$



Defining probability as a Measure over Possible worlds – infinite sets of variables, continuous random variables

$$\forall \omega \in \Omega, 0 \leq \mu(\omega), \int_{\omega} \mu(\omega) d\omega = 1, \quad P(f) = \int_{\omega|f} \mu(\omega) d\omega$$

When a random variable takes on real values the measure corresponds to a probability density function p . The probability that a random variable X takes values between a and b is given by

$$P(a \leq x \leq b) = \int_a^b p(x) dx$$

This definition can be generalized to handle vector valued random variables

Example:

$$p(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{(x-\mu)}{\sigma}\right)^2}$$

Note: we now have an infinite set of models



Inference by enumeration

- Start with the joint probability distribution:

	<i>toothache</i>		\neg <i>toothache</i>	
	<i>catch</i>	\neg <i>catch</i>	<i>catch</i>	\neg <i>catch</i>
<i>cavity</i>	.108	.012	.072	.008
\neg <i>cavity</i>	.016	.064	.144	.576

- For any proposition ϕ , sum the measures of atomic events where it is true: $P(\phi) = \sum_{\omega:\omega \models \phi} P(\omega)$



Inference by enumeration

- Start with the joint probability distribution:

	<i>toothache</i>		\neg <i>toothache</i>	
	<i>catch</i>	\neg <i>catch</i>	<i>catch</i>	\neg <i>catch</i>
<i>cavity</i>	.108	.012	.072	.008
\neg <i>cavity</i>	.016	.064	.144	.576

- For any proposition ϕ , sum the atomic events where it is true: $P(\phi) = \sum_{\omega: \omega \models \phi} P(\omega)$
- $P(\textit{toothache}) = 0.108 + 0.012 + 0.016 + 0.064 = 0.2$





Inference by enumeration

- Start with the joint probability distribution:

	<i>toothache</i>		\neg <i>toothache</i>	
	<i>catch</i>	\neg <i>catch</i>	<i>catch</i>	\neg <i>catch</i>
<i>cavity</i>	.108	.012	.072	.008
\neg <i>cavity</i>	.016	.064	.144	.576

- Can also compute conditional probabilities:

$$\begin{aligned}
 P(\neg \text{cavity} \mid \text{toothache}) &= \frac{P(\neg \text{cavity} \wedge \text{toothache})}{P(\text{toothache})} \\
 &= \frac{0.016 + 0.064}{0.108 + 0.012 + 0.016 + 0.064} \\
 &= 0.4
 \end{aligned}$$





Normalization

	<i>toothache</i>		\neg <i>toothache</i>	
	<i>catch</i>	\neg <i>catch</i>	<i>catch</i>	\neg <i>catch</i>
<i>cavity</i>	.108	.012	.072	.008
\neg <i>cavity</i>	.016	.064	.144	.576

- Denominator can be viewed as a **normalization constant** α
- $\mathbf{P}(Cavity \mid toothache) = \alpha \mathbf{P}(Cavity, toothache)$
 $= \alpha[\mathbf{P}(Cavity, toothache, catch) + \mathbf{P}(Cavity, toothache, \neg catch)]$
 $= \alpha[\langle 0.108, 0.016 \rangle + \langle 0.012, 0.064 \rangle]$
 $= \alpha \langle 0.12, 0.08 \rangle = \langle 0.6, 0.4 \rangle$
- General idea: compute distribution on query variable by fixing **evidence variables** and summing over **unobserved variables**





Inference by enumeration, continued

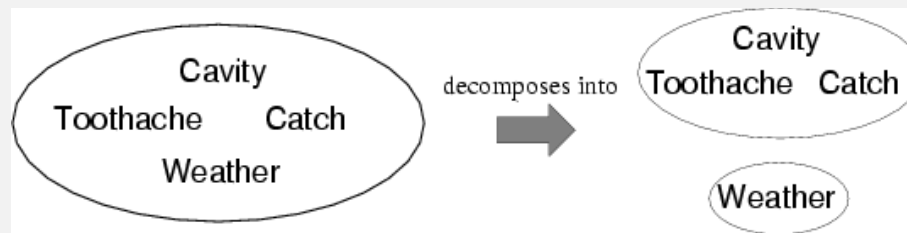
- Obvious problems:
 - Worst-case time complexity $O(d^n)$ where d is the largest arity
 - Space complexity $O(d^n)$ to store the joint distribution
 - How to find the numbers for $O(d^n)$ entries?



Independence

- A and B are independent iff

$$\mathbf{P}(A/B) = \mathbf{P}(A) \quad \text{or} \quad \mathbf{P}(B/A) = \mathbf{P}(B) \quad \text{or} \quad \mathbf{P}(A, B) = \mathbf{P}(A) \mathbf{P}(B)$$



$$\mathbf{P}(\textit{Toothache}, \textit{Catch}, \textit{Cavity}, \textit{Weather}) \\ = \mathbf{P}(\textit{Toothache}, \textit{Catch}, \textit{Cavity}) \mathbf{P}(\textit{Weather})$$

- 32 entries reduced to 12;
- n independent variables, $O(2^n)$ reduced to $O(n)$
- Absolute independence powerful but rare
- How can we manage a large numbers of variables?



Conditional independence

- $\mathbf{P}(\textit{Toothache}, \textit{Cavity}, \textit{Catch})$ has $2^3 - 1 = 7$ independent entries
- If I have a cavity, the probability that the probe catches in it doesn't depend on whether I have a toothache:
 - $\mathbf{P}(\textit{catch} \mid \textit{toothache}, \textit{cavity}) = \mathbf{P}(\textit{catch} \mid \textit{cavity})$
- The same independence holds if I haven't got a cavity:
 - $\mathbf{P}(\textit{catch} \mid \textit{toothache}, \neg \textit{cavity}) = \mathbf{P}(\textit{catch} \mid \neg \textit{cavity})$
- *Catch* is **conditionally independent** of *Toothache* given *Cavity*:
 - $\mathbf{P}(\textit{Catch} \mid \textit{Toothache}, \textit{Cavity}) = \mathbf{P}(\textit{Catch} \mid \textit{Cavity})$





Conditional independence

- *Catch* is **conditionally independent** of *Toothache* given *Cavity*:
 - $\mathbf{P}(\textit{Catch} \mid \textit{Toothache}, \textit{Cavity}) = \mathbf{P}(\textit{Catch} \mid \textit{Cavity})$
- Equivalent statements:
 - $\mathbf{P}(\textit{Toothache} \mid \textit{Catch}, \textit{Cavity}) = \mathbf{P}(\textit{Toothache} \mid \textit{Cavity})$
 - $\mathbf{P}(\textit{Toothache}, \textit{Catch} \mid \textit{Cavity}) = \mathbf{P}(\textit{Toothache} \mid \textit{Cavity}) \mathbf{P}(\textit{Catch} \mid \textit{Cavity})$

Conditional independence

- Write out full joint distribution using chain rule:

$$\mathbf{P}(\textit{Toothache}, \textit{Catch}, \textit{Cavity})$$

$$= \mathbf{P}(\textit{Toothache} \mid \textit{Catch}, \textit{Cavity}) \mathbf{P}(\textit{Catch}, \textit{Cavity})$$

$$= \mathbf{P}(\textit{Toothache} \mid \textit{Catch}, \textit{Cavity}) \mathbf{P}(\textit{Catch} \mid \textit{Cavity}) \mathbf{P}(\textit{Cavity})$$

$$= \mathbf{P}(\textit{Toothache} \mid \textit{Cavity}) \mathbf{P}(\textit{Catch} \mid \textit{Cavity}) \mathbf{P}(\textit{Cavity})$$

i.e., $2 + 2 + 1 = 5$ independent numbers

- **Conditional independence**
 - often reduces the size of the representation of the joint distribution from exponential in n to linear in n
 - Is one of the most basic and robust form of knowledge about uncertain environments



Conditional Independence

- X is **conditionally independent** of Y given Z (written $I(X,Z,Y)$) if the probability distribution governing X is independent of the value of Y given the value of Z :
- $P(X | Y, Z) = P(X | Z)$ that is,

$$(\forall x_i, y_j, z_k) P(X = x_i | Y = y_j, Z = z_k) = P(X = x_i | Z = z_k)$$





Independence is symmetric: $I(X, Y, Z) = I(Z, Y, X)$

- Assume: $P(X|Y, Z) = P(X|Y)$
- X and Z are independent given Y

$$P(Z | X, Y) = \frac{P(X, Y | Z)P(Z)}{P(X, Y)} \quad (\text{Bayes' s Rule})$$

$$= \frac{P(Y | Z)P(X | Y, Z)P(Z)}{P(X | Y)P(Y)} \quad (\text{Chain Rule})$$

$$= \frac{P(Y | Z)P(X | Y)P(Z)}{P(X | Y)P(Y)} \quad (\text{By Assumption})$$

$$= \frac{P(Y | Z)P(Z)}{P(Y)} = P(Z | Y) \quad (\text{Bayes' s Rule})$$





Bayes Rule

Does patient have cancer or not?

A patient takes a lab test and the result comes back positive. The test returns a correct positive result in only 98% of the cases in which the disease is actually present, and a correct negative result in only 97% of the cases in which the disease is not present. Furthermore, .008 of the entire population have this cancer.

$$P(cancer) =$$

$$P(\neg cancer) =$$

$$P(+ | cancer) =$$

$$P(- | cancer) =$$

$$P(+ | \neg cancer) =$$

$$P(- | \neg cancer) =$$



Bayes Rule

Does patient have cancer or not?

$$P(\text{cancer}) = 0.008$$

$$P(\neg\text{cancer}) = 0.992$$

$$P(+ | \text{cancer}) = 0.98$$

$$P(- | \text{cancer}) = 0.02$$

$$P(+ | \neg\text{cancer}) = 0.03$$

$$P(- | \neg\text{cancer}) = 0.97$$

$$P(\text{cancer} | +) = \frac{P(+ | \text{cancer})P(\text{cancer})}{P(+)};$$

$$P(\neg\text{cancer} | +) = \frac{P(+ | \neg\text{cancer})P(\neg\text{cancer})}{P(+)}$$

$$P(\text{cancer} | +)P(+) = 0.98 \times 0.008 = 0.0078;$$

$$P(\neg\text{cancer} | +)P(+) = 0.03 \times 0.992 = 0.0298$$

$$P(+) = 0.0078 + 0.0298$$

$$P(\text{cancer} | +) = 0.21;$$

$$P(\neg\text{cancer} | +) = 0.79$$

The patient, more likely than not, does not have cancer





Bayes Rule

- Product rule
 - $P(a \wedge b) = P(a | b) P(b) = P(b | a) P(a)$
 - **Bayes' rule:** $P(a | b) = P(b | a) P(a) / P(b)$

- In distribution form

$$P(Y|X) = P(X|Y) P(Y) / P(X) = \alpha P(X|Y) P(Y)$$





Probabilistic KR: The story so far

- Probability is a rigorous formalism for uncertain knowledge
- **Joint probability distribution specifies probability of every atomic event**
- Queries can be answered by summing over atomic events
- **Independence** and **conditional independence** provide the basis for compact representation of joint probability distributions
- Graph theory provides a basis for efficient computation
-



Building Probabilistic Models – Conditional Independence

- Random variable X is **conditionally independent** of Y given Z if the probability distribution governing X is independent of the value of Y given the value of Z :
- $P(X | Y, Z) = P(X | Z)$ that is, if

$$(\forall x_i, y_j, z_k) P(X = x_i | Y = y_j, Z = z_k) = P(X = x_i | Z = z_k)$$



Conditional Independence

$$\begin{aligned}P(\textit{Thunder} = 1 \mid \textit{Rain} = 1, \textit{Lightning} = 1) &= P(\textit{Thunder} = 1 \mid \textit{Lightening} = 1) \\ &= P(\textit{Thunder} = 1 \mid \textit{Rain} = 0, \textit{Lightening} = 1)\end{aligned}$$

$$\begin{aligned}P(\textit{Thunder} = 1 \mid \textit{Rain} = 1, \textit{Lightning} = 0) &= P(\textit{Thunder} = 1 \mid \textit{Lightening} = 0) \\ &= P(\textit{Thunder} = 1 \mid \textit{Rain} = 0, \textit{Lightening} = 0)\end{aligned}$$

$$\begin{aligned}P(\textit{Thunder} = 0 \mid \textit{Rain} = 1, \textit{Lightning} = 1) &= P(\textit{Thunder} = 0 \mid \textit{Lightening} = 1) \\ &= P(\textit{Thunder} = 0 \mid \textit{Rain} = 0, \textit{Lightening} = 1)\end{aligned}$$

$$\begin{aligned}P(\textit{Thunder} = 0 \mid \textit{Rain} = 1, \textit{Lightning} = 0) &= P(\textit{Thunder} = 0 \mid \textit{Lightening} = 0) \\ &= P(\textit{Thunder} = 0 \mid \textit{Rain} = 0, \textit{Lightening} = 0)\end{aligned}$$





Conditional Independence

Let Z_1, \dots, Z_n and W be random variables
on a given event space.

Z_1, \dots, Z_n are mutually independent given W if

$$P(Z_1, Z_2, \dots, Z_n | W) = \prod_{i=1}^n P(Z_i | W)$$

$P(Z_1 | Z_2, W) = P(Z_1 | W)$ if Z_1 and Z_2 are independent.

Note that these represent sets of equations, for all
possible value assignments to random variables

Independence Properties of Random Variables

Let $\mathbf{W}, \mathbf{X}, \mathbf{Y}, \mathbf{Z}$ be pairwise disjoint sets of random variables on a given event space.

Let $I(\mathbf{X}, \mathbf{Y}, \mathbf{Z})$ denote that \mathbf{X} and \mathbf{Z} are independent given \mathbf{Y} .

That is, $P(\mathbf{X} \cup \mathbf{Z} | \mathbf{Y}) = P(\mathbf{X} | \mathbf{Y})P(\mathbf{Z} | \mathbf{Y})$, or $P(\mathbf{X} | \mathbf{Y} \cup \mathbf{Z}) = P(\mathbf{X} | \mathbf{Y})$. Then :

a. $I(\mathbf{X}, \mathbf{Z}, \mathbf{Y}) \Rightarrow I(\mathbf{Y}, \mathbf{Z}, \mathbf{X})$

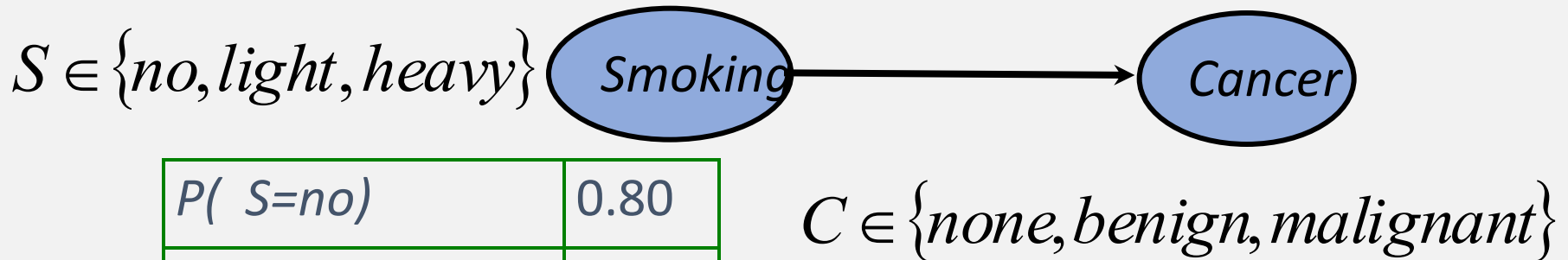
b. $I(\mathbf{X}, \mathbf{Z}, \mathbf{Y} \cup \mathbf{W}) \Rightarrow I(\mathbf{X}, \mathbf{Z}, \mathbf{Y})$

c. $I(\mathbf{X}, \mathbf{Z}, \mathbf{Y} \cup \mathbf{W}) \Rightarrow I(\mathbf{X}, \mathbf{Z} \cup \mathbf{W}, \mathbf{Y})$

d. $I(\mathbf{X}, \mathbf{Z}, \mathbf{Y}) \wedge I(\mathbf{X}, \mathbf{Z} \cup \mathbf{Y}, \mathbf{W}) \Rightarrow I(\mathbf{X}, \mathbf{Z}, \mathbf{Y} \cup \mathbf{W})$

Proof : Follows from definition of independence.

Bayesian Networks



$P(S=no)$	0.80
$P(S=light)$	0.15
$P(S=heavy)$	0.05

$Smoking=$	<i>no</i>	<i>light</i>	<i>heavy</i>
$P(C=none)$	0.96	0.88	0.60
$P(C=benign)$	0.03	0.08	0.25
$P(C=malig)$	0.01	0.04	0.15



Product Rule

- $P(C,S) = P(C|S) P(S)$

$S \Downarrow$	$C \Rightarrow$	<i>none</i>	<i>benign</i>	<i>malignant</i>
<i>no</i>		0.768	0.024	0.008
<i>light</i>		0.132	0.012	0.006
<i>heavy</i>		0.035	0.010	0.005

Marginalization

$S \downarrow C \Rightarrow$	<i>none</i>	<i>benign</i>	<i>malig</i>	total
<i>no</i>	0.768	0.024	0.008	.80
<i>light</i>	0.132	0.012	0.006	.15
<i>heavy</i>	0.035	0.010	0.005	.05
total	0.935	0.046	0.019	

} $P(\text{Smoke})$

} $P(\text{Cancer})$



Bayes Rule Revisited

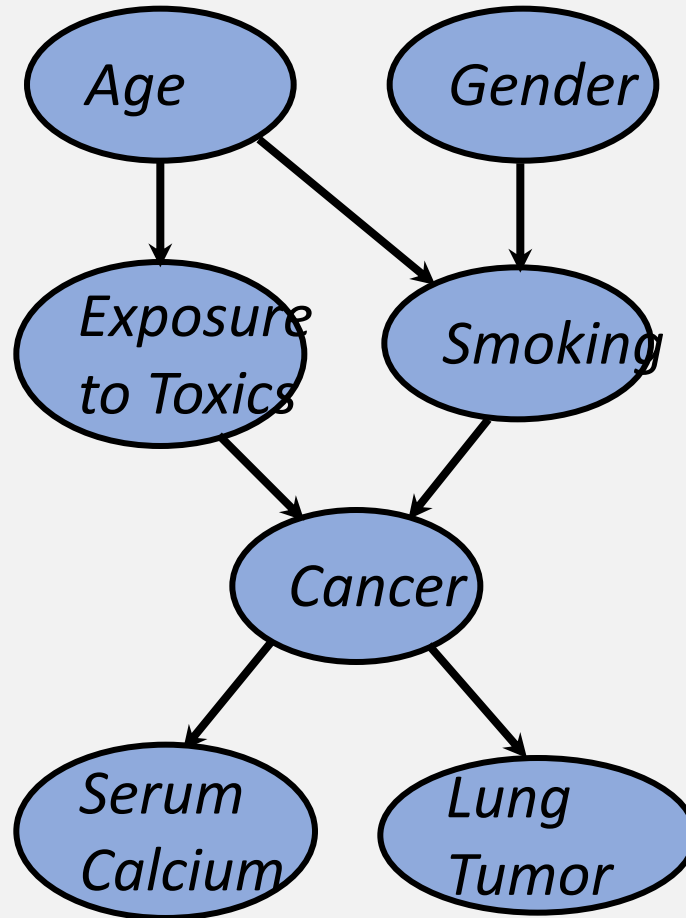
$$P(S | C) = \frac{P(C | S)P(S)}{P(C)} = \frac{P(C, S)}{P(C)}$$

$S \Downarrow$ $C \Rightarrow$	<i>none</i>	<i>benign</i>	<i>malig</i>
<i>no</i>	0.768/.935	0.024/.046	0.008/.019
<i>light</i>	0.132/.935	0.012/.046	0.006/.019
<i>heavy</i>	0.030/.935	0.015/.046	0.005/.019

<i>Cancer=</i>	<i>none</i>	<i>benign</i>	<i>malignant</i>
$P(S=no)$	0.821	0.522	0.421
$P(S=light)$	0.141	0.261	0.316
$P(S=heavy)$	0.037	0.217	0.263

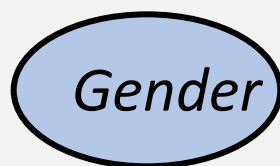
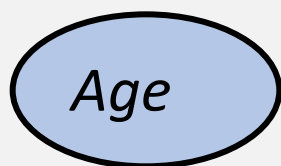


A Bayesian Network





Independence



Age and Gender are independent.

$$P(A, G) = P(G)P(A)$$

$$P(A|G) = P(A) \quad A \perp G$$

$$P(G|A) = P(G) \quad G \perp A$$

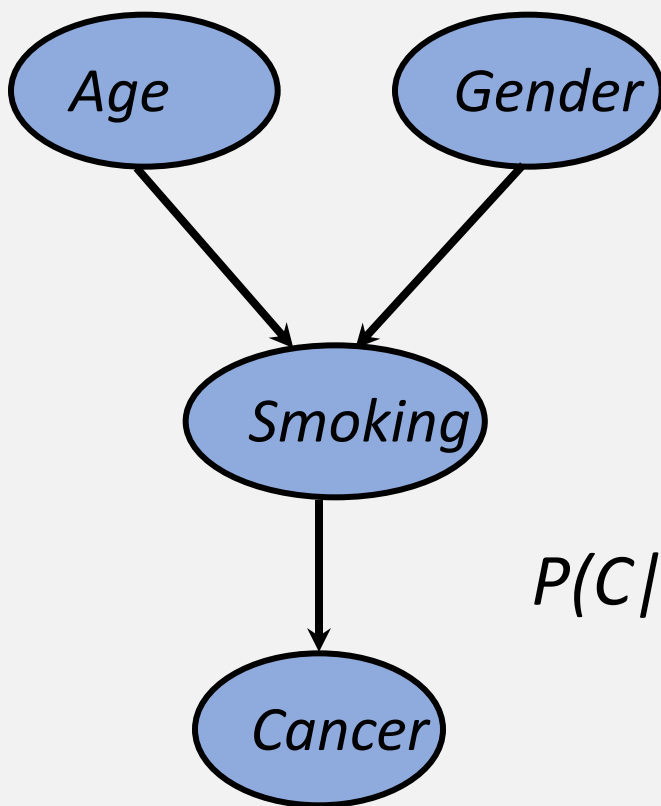
$$P(A, G) = P(G|A) P(A) = P(G)P(A)$$

$$P(A, G) = P(A|G) P(G) = P(A)P(G)$$





Conditional Independence



Cancer is independent
of *Age* and *Gender*
given *Smoking*.

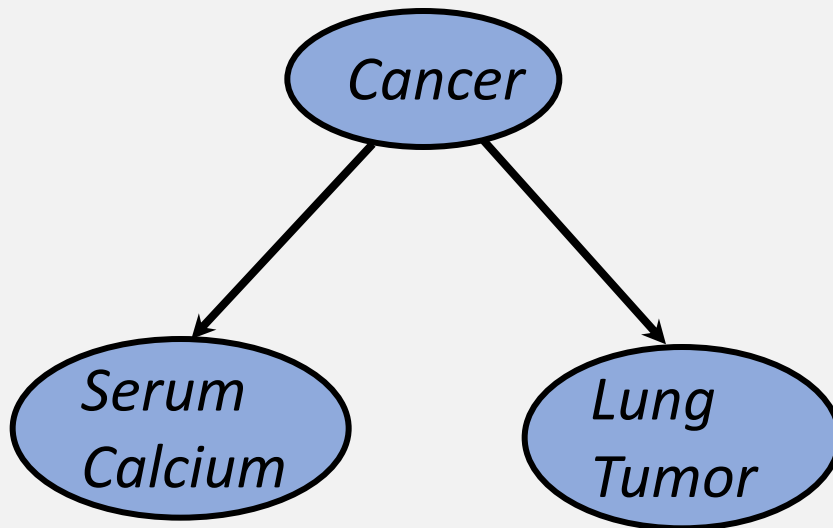
$$P(C|A,G,S) = P(C|S) \quad C \perp A,G \mid S$$



More Conditional Independence: Naïve Bayes

Serum Calcium and Lung Tumor are dependent

Serum Calcium is independent of Lung Tumor, given Cancer



$$P(L|SC, C) = P(L|C)$$



Probabilistic Graphical Models

- The Probabilistic graphical models e.g., Bayes networks, explicitly model conditional independence among subsets of variables to yield a graphical representation of probability distributions that admit such independence

$$P(X_1, X_2, \dots, X_n) = \prod_{i=1}^n P(X_i \mid \mathbf{Pa}_i)$$

$$\mathbf{Pa}_i = \text{parents}(X_i)$$



Bayesian network

- Bayesian network is a directed acyclic graph (DAG) in which the nodes represent random variables
- Each node is annotated with a probability distribution $P(X_i / Parents(X_i))$ representing the dependency of that node on its parents in the DAG
- Each node is asserted to be conditionally independent of its non-descendants, given its immediate predecessors
- Arcs represent direct dependencies

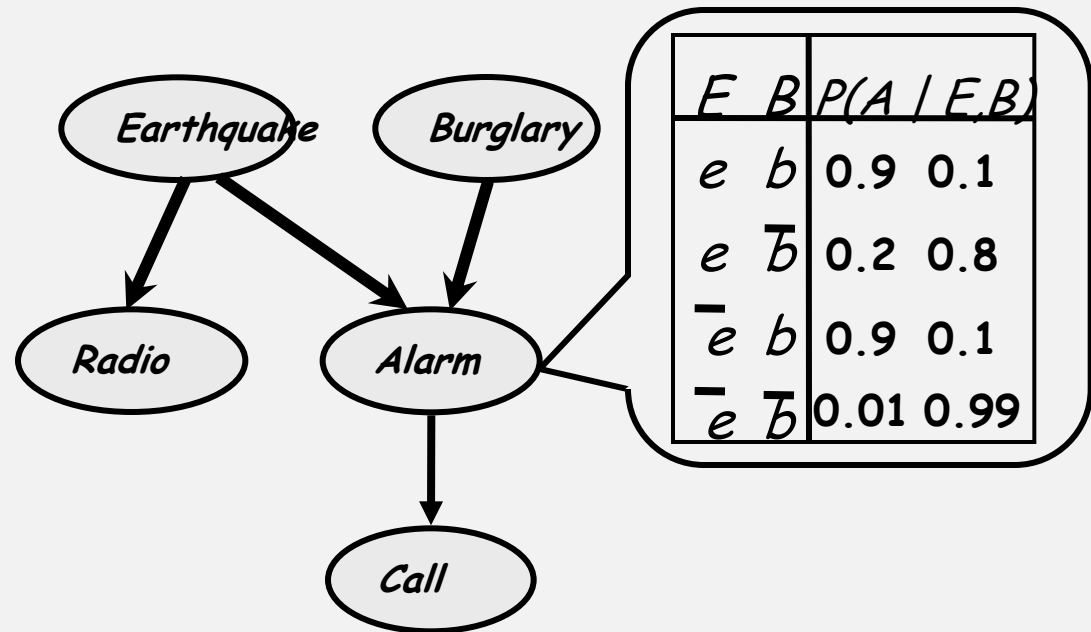


Conditional Independence

- X is **conditionally independent** of Y given Z if the probability distribution governing X is independent of the value of Y given the value of Z :
- $P(X | Y, Z) = P(X | Z)$ that is,

$$(\forall x_i, y_j, z_k) P(X = x_i | Y = y_j, Z = z_k) = P(X = x_i | Z = z_k)$$

Bayesian Networks



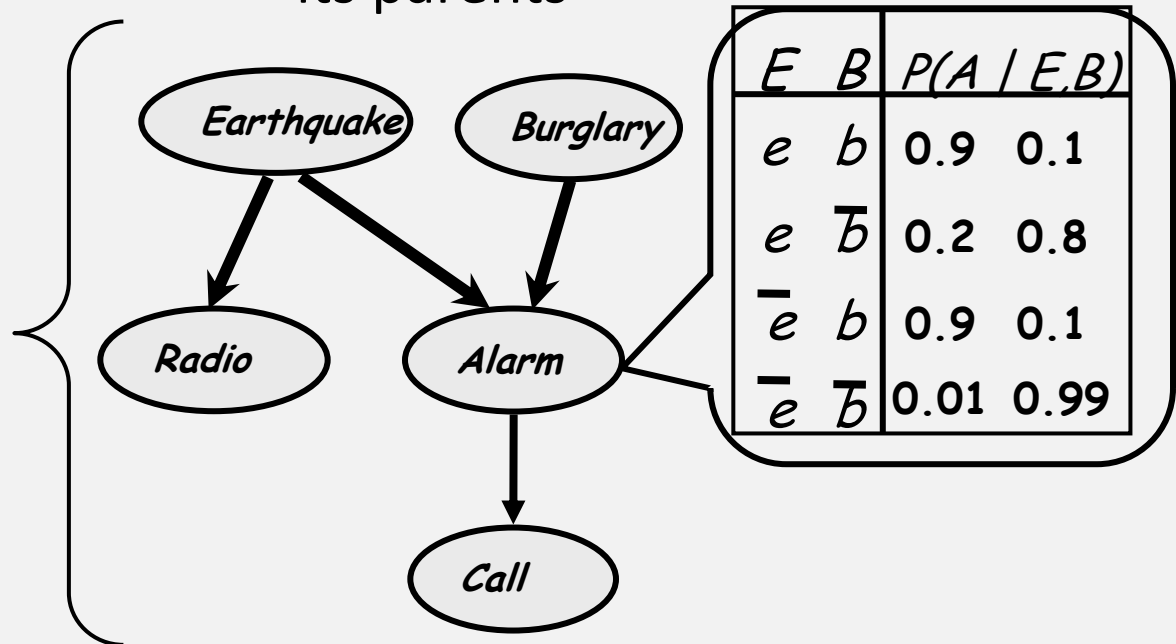


Bayesian Networks

- **Qualitative part**
statistical independence
statements represented
in the form of a directed
acyclic graph
(DAG)
 - Nodes - random variables
 - Edges – direct influence

Quantitative part

Conditional probability
distributions – one for each
random variable conditioned on
its parents

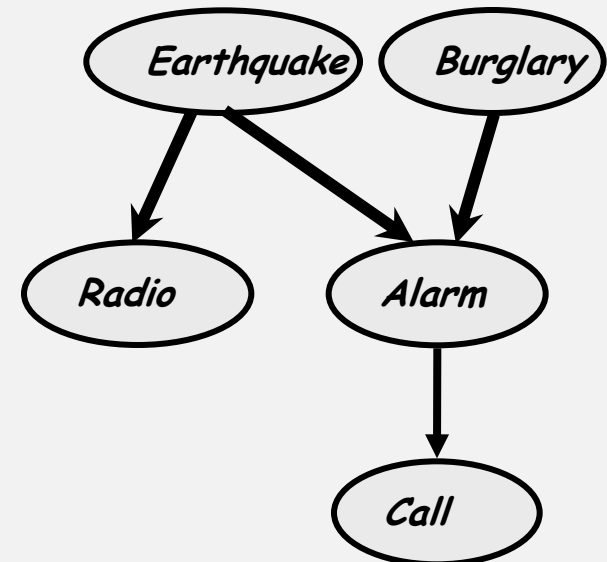


Efficient factorized representation of probability distributions via conditional independence

- Nodes are independent of non-descendants given their parents

d-separation:

- a graph theoretic criterion for checking implicit independence assertions
- can be computed in linear time (in the number of edges)



What independences does a Bayes Net model?

- In order for a Bayesian network to model a probability distribution, the following must be true by definition:
- Each variable is conditionally independent of all its non-descendants in the graph given the value of all its parents.

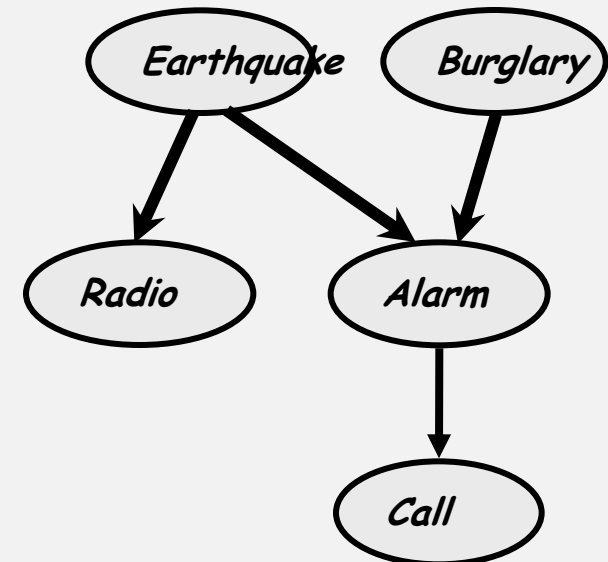
This implies

$$P(X_1 \dots X_n) = \prod_{i=1}^n P(X_i \mid \text{parents}(X_i))$$

$$P(E, B, R, A, C) =$$

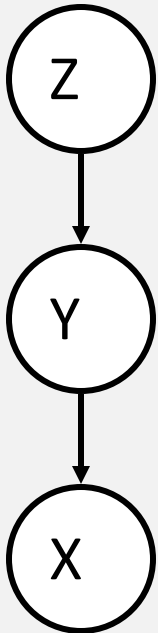
$$P(E)P(B)P(R \mid E)P(A \mid E, B)P(C \mid A)$$

but what else does it imply?



What Independences does a Bayes Network model?

Example:



Given Y , does learning the value of Z tell us nothing new about X ?

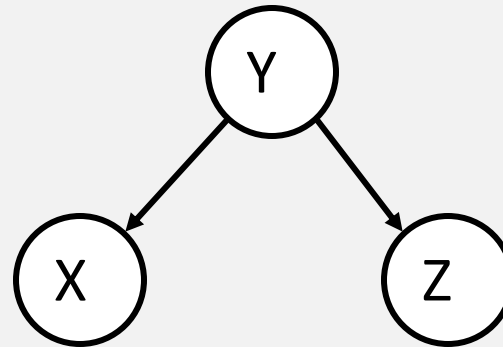
i.e., is $P(X|Y, Z)$ equal to $P(X | Y)$?

Yes. Since we know the value of all of X 's parents (namely, Y), and Z is not a descendant of X , X is conditionally independent of Z .

Also, since independence is symmetric, $P(Z|Y, X) = P(Z|Y)$.

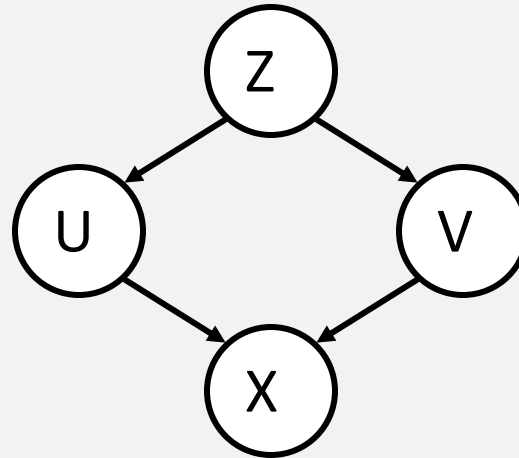
What Independences does a Bayes Network model?

- Let $I(X,Y,Z)$ represent X and Z being conditionally independent given Y .



- $I(X,Y,Z)$? Yes, just as in previous example: All X 's parents given, and Z is not a descendant.

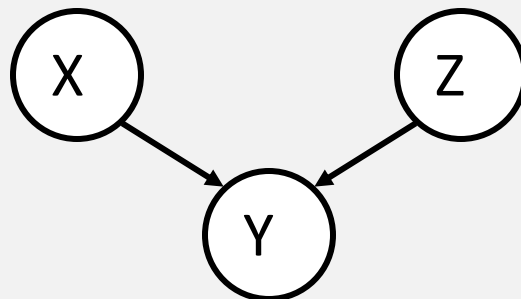
What Independences does a Bayes Network model?



- $I(X, \{U\}, Z)$? No.
- $I(X, \{U, V\}, Z)$? Yes.



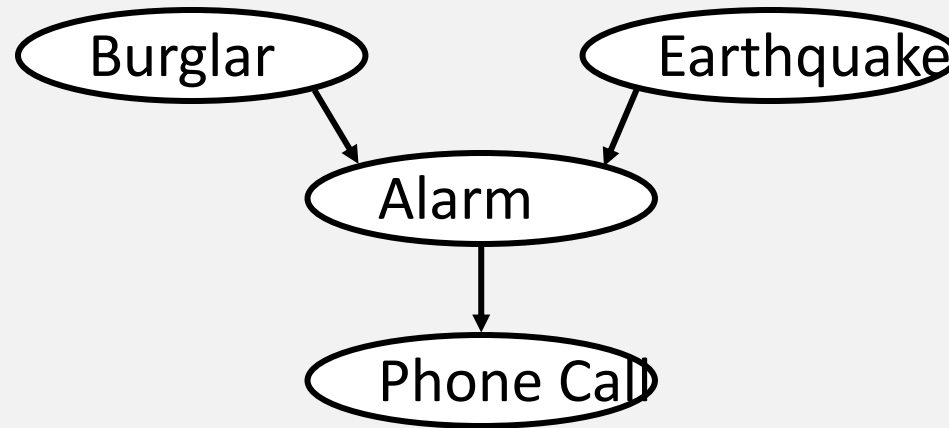
Dependency induced by V-structures



- X has no parents, so we know all its parents' values trivially
- Z is not a descendant of X
- So, $I(X, \{\}, Z)$, even though there is an undirected path from X to Z through an unknown variable Y .
- What if we do know the value of Y ? Or one of its descendants?

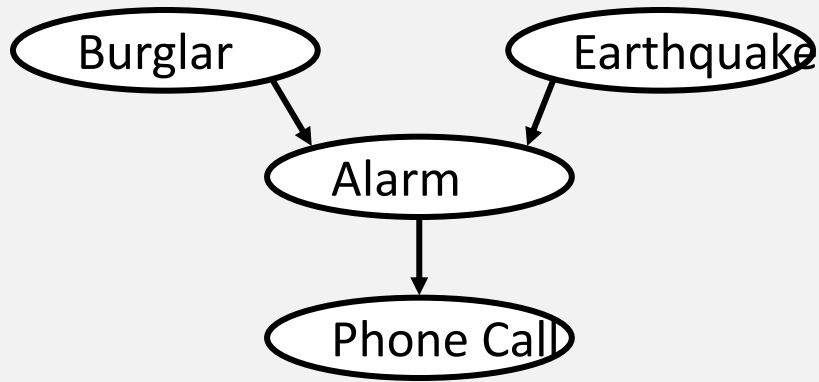


The Burglar Alarm example



- Your house has a twitchy burglar alarm that is also sometimes triggered by earthquakes.
- Earth arguably doesn't care whether your house is currently being burgled
- While you are on vacation, one of your neighbors calls and tells you your home's burglar alarm is ringing.





- But now suppose you learn that there was a medium-sized earthquake in your neighborhood. ...Probably not a burglar after all.
- Earthquake “explains away” the hypothetical burglar.
- But then it must NOT be the case that $I(\text{Burglar}, \{\text{Phone Call}\}, \text{Earthquake})$, even though $I(\text{Burglar}, \{\}, \text{Earthquake})!$



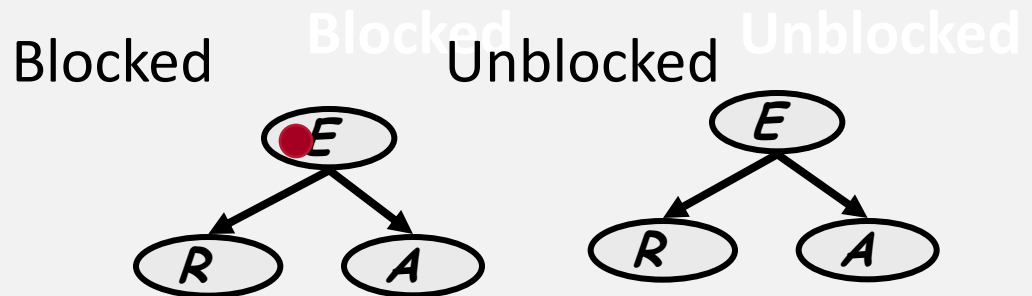
d-separation

- Fortunately, there is a relatively simple algorithm for determining whether two variables in a Bayesian network are conditionally independent given some other variables:
 - *d-separation*.
- Two variables are independent if all paths between them are blocked by evidence
- Three cases:
 - Common cause
 - Intermediate cause
 - Common Effect

d-separation

- Two variables are independent if all paths between them are blocked by evidence
- Three cases:
 - Common cause
 - Intermediate cause
 - Common Effect

Evidence may be transmitted through a diverging connection unless it is instantiated.



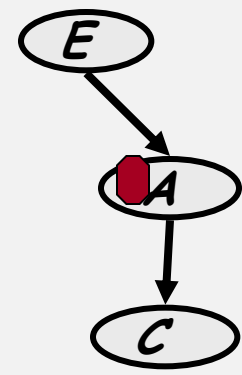
- If we do not know whether an earthquake occurred, then radio announcement can influence our belief about the alarm having gone off.
- If we know that earthquake occurred, then radio announcement gives no information about the alarm

d-separation

- Common cause
- Intermediate cause
- Common Effect

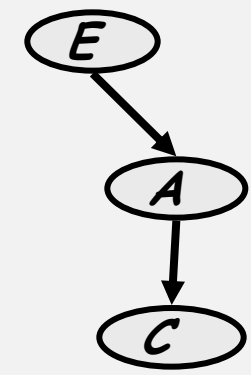
Blocked

Blocked



Unblocked

Unblocked

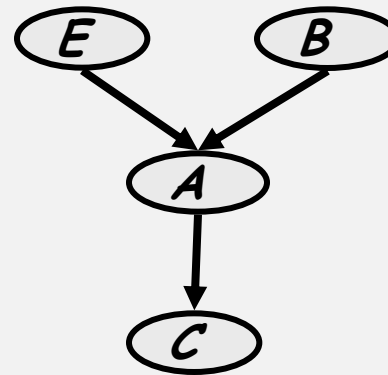


Information may be transmitted through a serial connection unless it is blocked (value set)

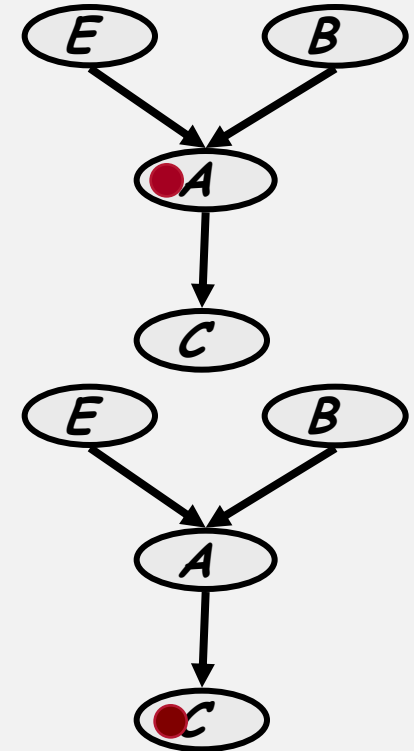
d-separation

- Common cause
- Intermediate cause
- Common Effect

Blocked



Unblocked



Information may be transmitted through a converging connection only if either the variable or one of its descendants has been set



d-separation

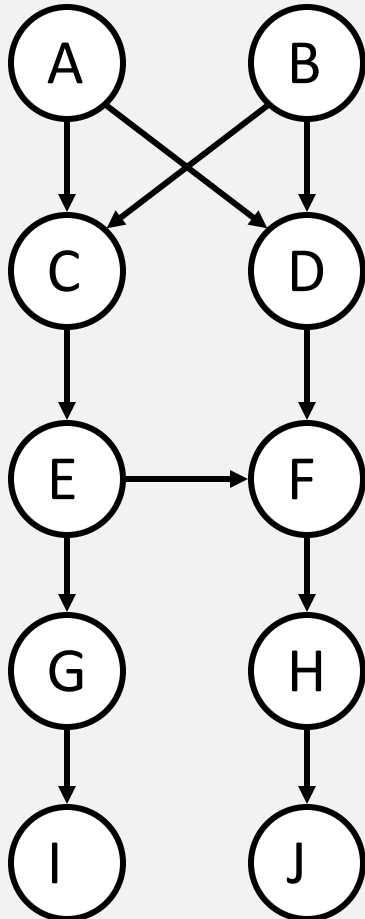
- **Definition:** X and Z are *d-separated* by a set of evidence variables E iff every undirected path from X to Z is “blocked” by evidence E



d-separation

- Theorem [Verma & Pearl, 1998]: If a set of evidence variables E d -separates X and Z in a Bayesian network's graph, then $I(X, E, Z)$.
- d -separation can be computed in linear time using a depth-first search like algorithm.
- We now have a fast algorithm for automatically inferring whether finding out about the value of one variable might give us any additional hints about some other variable, given what we already know.
- d -separation of X and Z by E is sufficient for asserting $I(X, E, Z)$, but not necessary.
 - Variables may actually be independent when they are not d -separated, depending on the actual probabilities involved

d-separation



$I(C, \{\}, D)?$

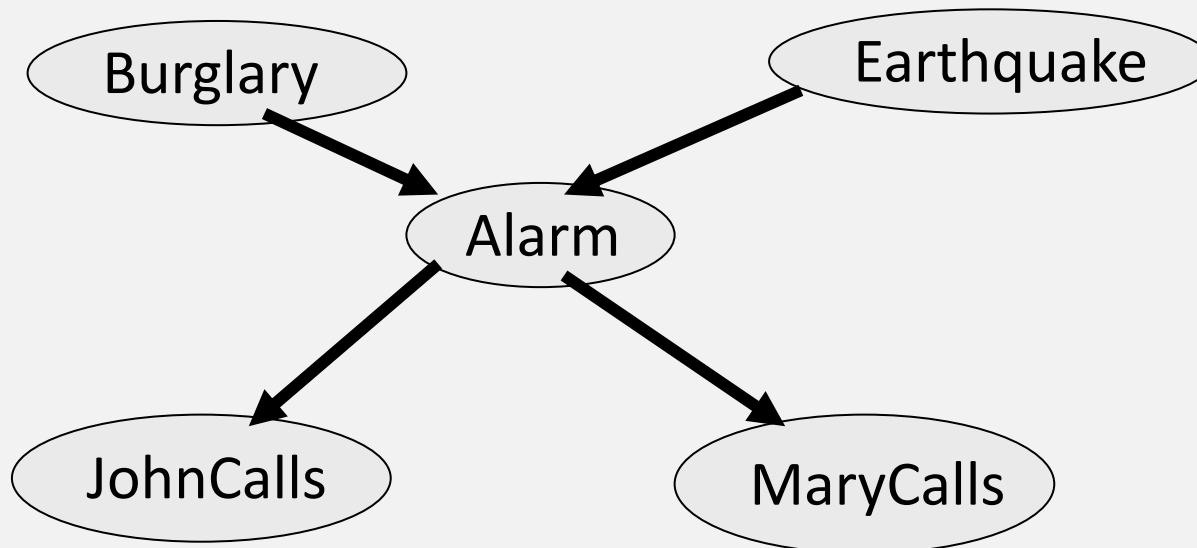
$I(C, \{A\}, D)?$

$I(C, \{A, B\}, D)?$

$I(C, \{A, B, J\}, D)?$

Markov Blanket

- A node is conditionally independent of all other nodes in the network given its parents, children, and children's parents -



Burglary is independent of John Calls and Mary Calls given Alarm and Earth Quake



Bayesian Networks: Summary

- Bayesian networks offer an efficient representation of probability distributions
- Efficient:
 - Local models
 - Independence (d -separation)
- Effective: Algorithms take advantage of structure to
 - Compute posterior probabilities
 - Compute most probable instantiation
 - Decision making

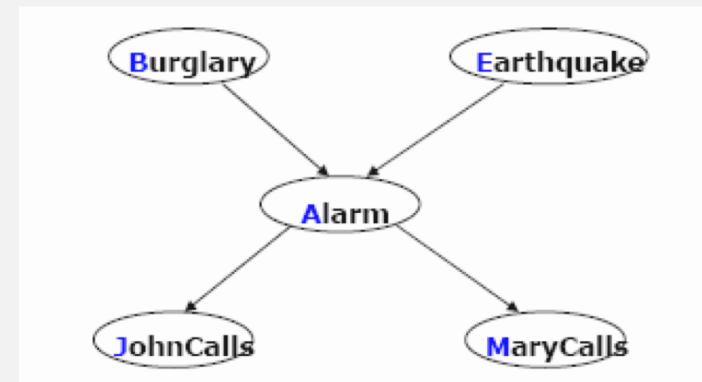
Inference in Bayesian network

Bad news:

- – Exact inference problem in BNs is NP-hard (Cooper)
- – Approximate inference is NP-hard (Dagum, Luby)

In practice, things are not so bad

- Exact inference
 - Inference in Simple Chains
 - Variable elimination
 - Clustering / join tree algorithms
- Approximate inference
 - Stochastic simulation / sampling methods
 - Markov chain Monte Carlo methods
 - Mean field theory





Computing joint probability distributions using a Bayesian network

- Any entry in the joint probability distribution can be calculated from the Bayesian network.
- We're just using the chain rule and conditional independence.

$$\begin{aligned}P(J, M, A, \neg B, \neg E) &= P(J \mid M, A, \neg B, \neg E)P(M, A, \neg B, \neg E) \\ &= P(J \mid A)P(M \mid A, \neg B, \neg E)P(A, \neg B, \neg E) \\ &= P(J \mid A)P(M \mid A)P(A \mid \neg B, \neg E)P(\neg B, \neg E) \\ &= P(J \mid A)P(M \mid A)P(A \mid \neg B, \neg E)P(\neg B)P(\neg E)\end{aligned}$$



Computing joint probabilities

General formula:

$$P(X_1, \dots, X_n) = P(X_1) \prod_{i=2}^n P(X_i \mid \text{Parents}(X_i))$$

- Joint distribution can be used to answer any query about the domain.
- Bayesian network represents the joint distribution
- Any query about the domain can be answered using a BN
- Tradeoff: A BN can be much more concise, but you need to calculate, rather than look up in a table, probabilities from the joint distribution





Inference in Bayesian Networks

- Bayesian networks are a compact encoding of the full joint probability distribution over N variables that makes conditional independence assumptions between these variables explicit.
- We can use Bayesian networks to compute any probability of interest over the given variables.
- Now we look at Inference in more detail



Inference in Bayesian Networks

Find $P(Q=q|E=e)$

- Q the query variable(s)
- E set of evidence variables

$$P(q|e) = P(q,e) / P(e)$$

X_1, \dots, X_n are network variables except Q, E

$$P(q,e) = \sum_{x_1, x_2, \dots, x_n} P(q, e, X_1, X_2, \dots, X_n)$$

Basic Inference

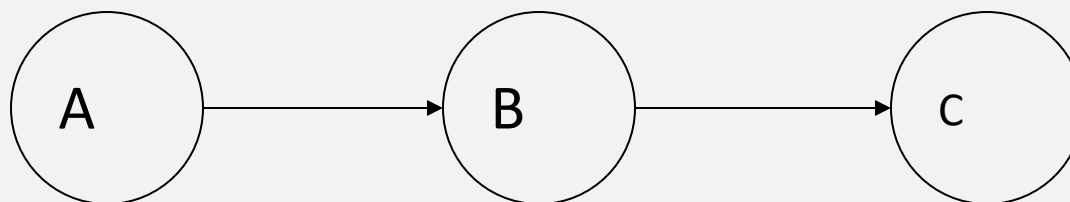


$$P(b) = ?$$

$$P(b) = \sum_a P(a, b) = \sum_a P(b|a) P(a)$$



Basic Inference



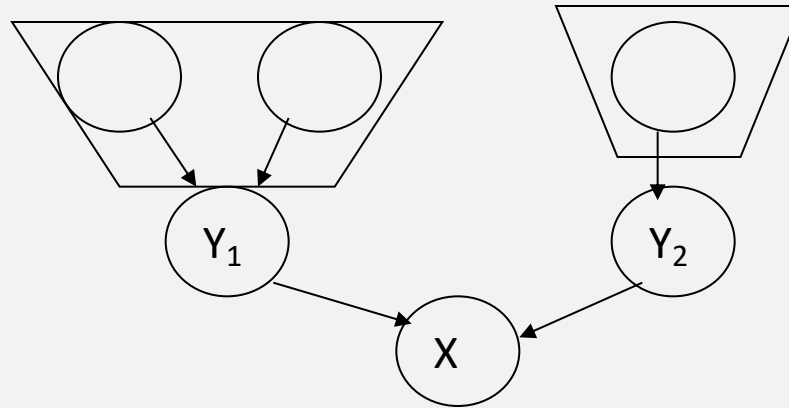
$$P(b) = \sum_a P(a, b) = \sum_a P(b | a) P(a)$$

$$P(c) = \sum_b P(c | b) P(b)$$

$$\begin{aligned} P(c) &= \sum_{a,b} P(a, b, c) = \sum_{a,b} P(c | b, a) P(b | a) P(a) \\ &= \sum_{a,b} P(c | b) P(b | a) P(a) \\ &= \sum_{a,b} P(c | b) P(b) \end{aligned}$$



Inference in trees

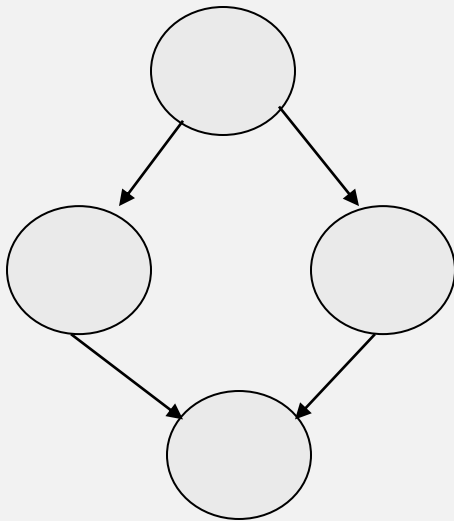


$$P(X) = \sum_{y_1, y_2} P(X, Y_1, Y_2) = \sum_{y_1, y_2} P(X | Y_1, Y_2) P(Y_1, Y_2) = \sum_{y_1, y_2} P(X | Y_1, Y_2) P(Y_1) P(Y_2)$$

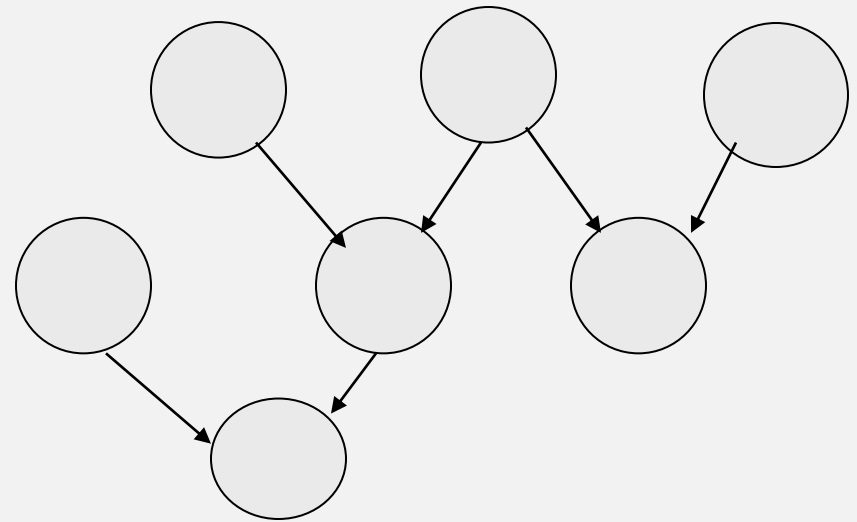


Polytrees

- A network is *singly connected* (a *polytree*) if it contains no **undirected loops**.



Not a polytree



Polytree



Inference in polytrees

- **Theorem:** Inference in polytrees can be performed in time that is polynomial in the number of variables.
- **Main idea:** in variable elimination, need only maintain distributions over single nodes at any step.



Inference with Bayesian Networks

- Inference in polytrees can be performed efficiently
- Inference with DAG is NP-Hard
 - Proof by reduction of SAT to Bayesian network inference



Approaches to inference

- Exact inference
 - Inference in Simple Chains
 - Variable elimination
 - Clustering / join tree algorithms
- Approximate inference
 - Stochastic simulation / sampling methods
 - Markov chain Monte Carlo methods
 - Mean field theory

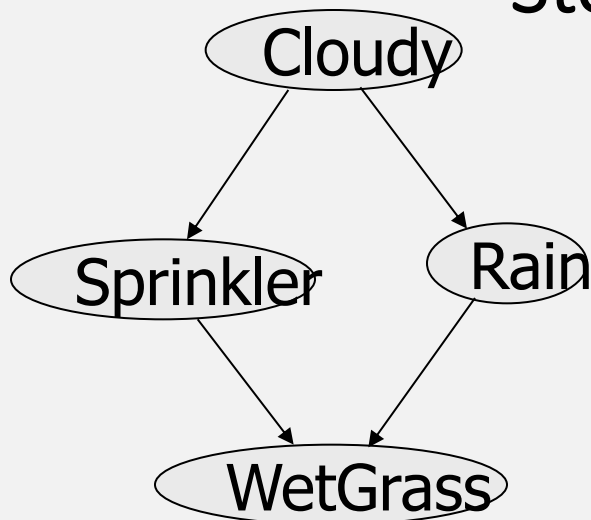


Approximate Inference: Stochastic simulation

- Exact inference in Bayesian Networks is hard
- Suppose you are given values for some subset of the variables, G , and want to infer values for unknown variables, U
- Randomly generate a very large number of instantiations from the BN
 - Generate instantiations for **all** variables – start at root variables and work your way “forward”
- Only keep those instantiations that are consistent with the values for G
- Use the frequency of values for U to get estimated probabilities
- Accuracy of the results depends on the size of the sample (asymptotically approaches exact results)



Stochastic Simulation



$P(\text{WetGrass} | \text{Cloudy})?$

$$P(\text{WetGrass} | \text{Cloudy}) \\ = P(\text{WetGrass}, \text{Cloudy}) / P(\text{Cloudy})$$

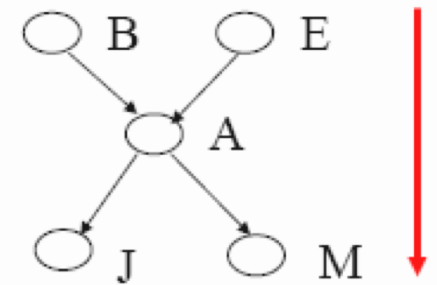
1. Draw N samples from the BN by repeating 1.1 and 1.2
 - 1.1. Guess Cloudy at random according to $P(\text{Cloudy})$
 - 1.2. For each guess of Cloudy, guess Sprinkler and Rain, then WetGrass
2. Compute the ratio of the # runs where WetGrass and Cloudy are True over the # runs where Cloudy is True

Stochastic simulation

- The probability is approximated using sample frequencies

BN sampling:

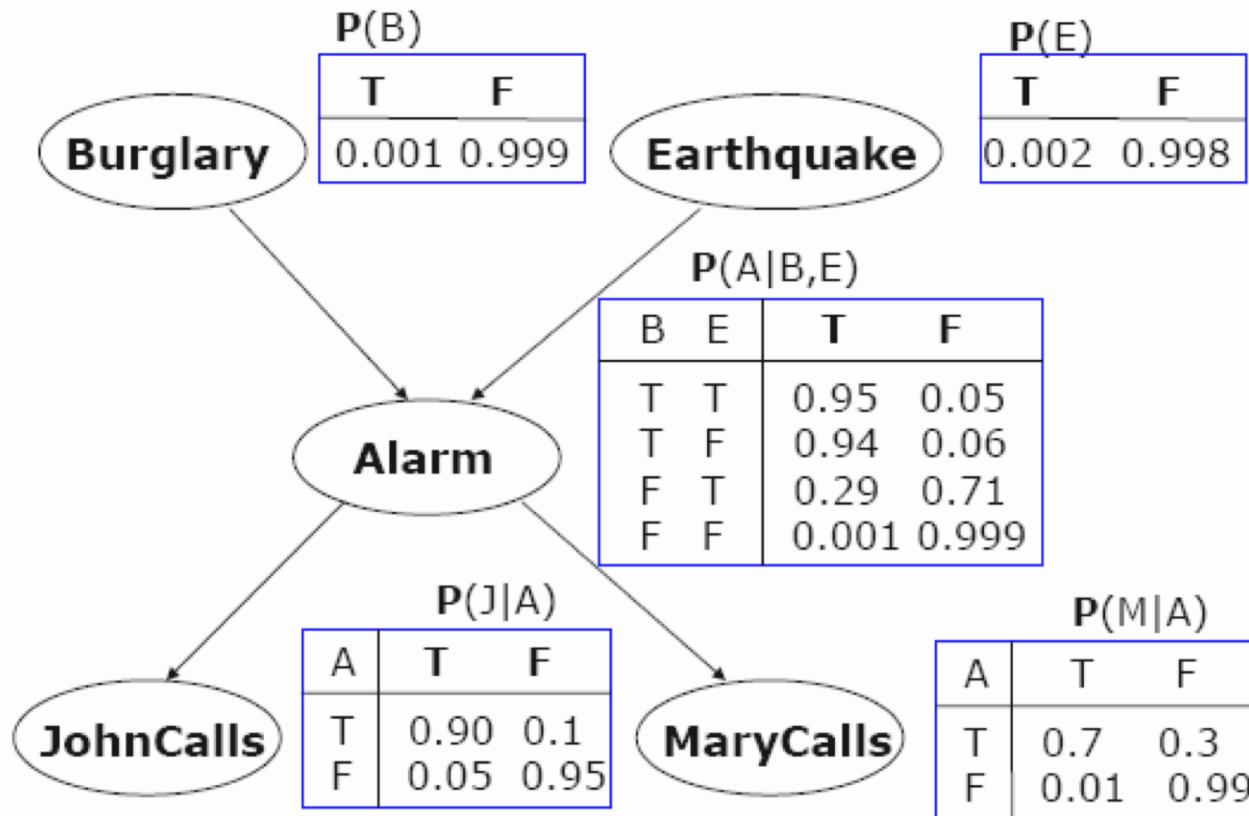
- Generate sample in a top down manner, following the links in BN
- A sample is an assignment of values to all variables



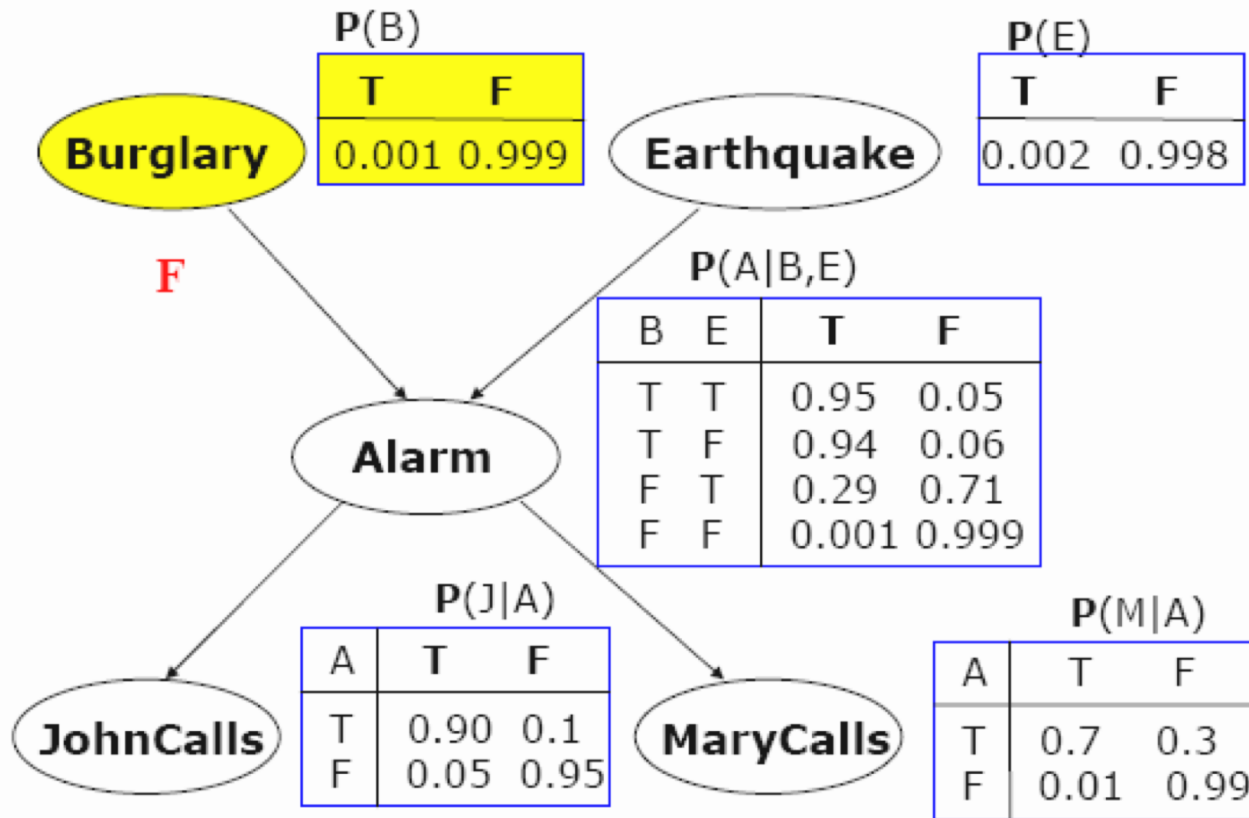
BN Sampling Example

Goal: To infer

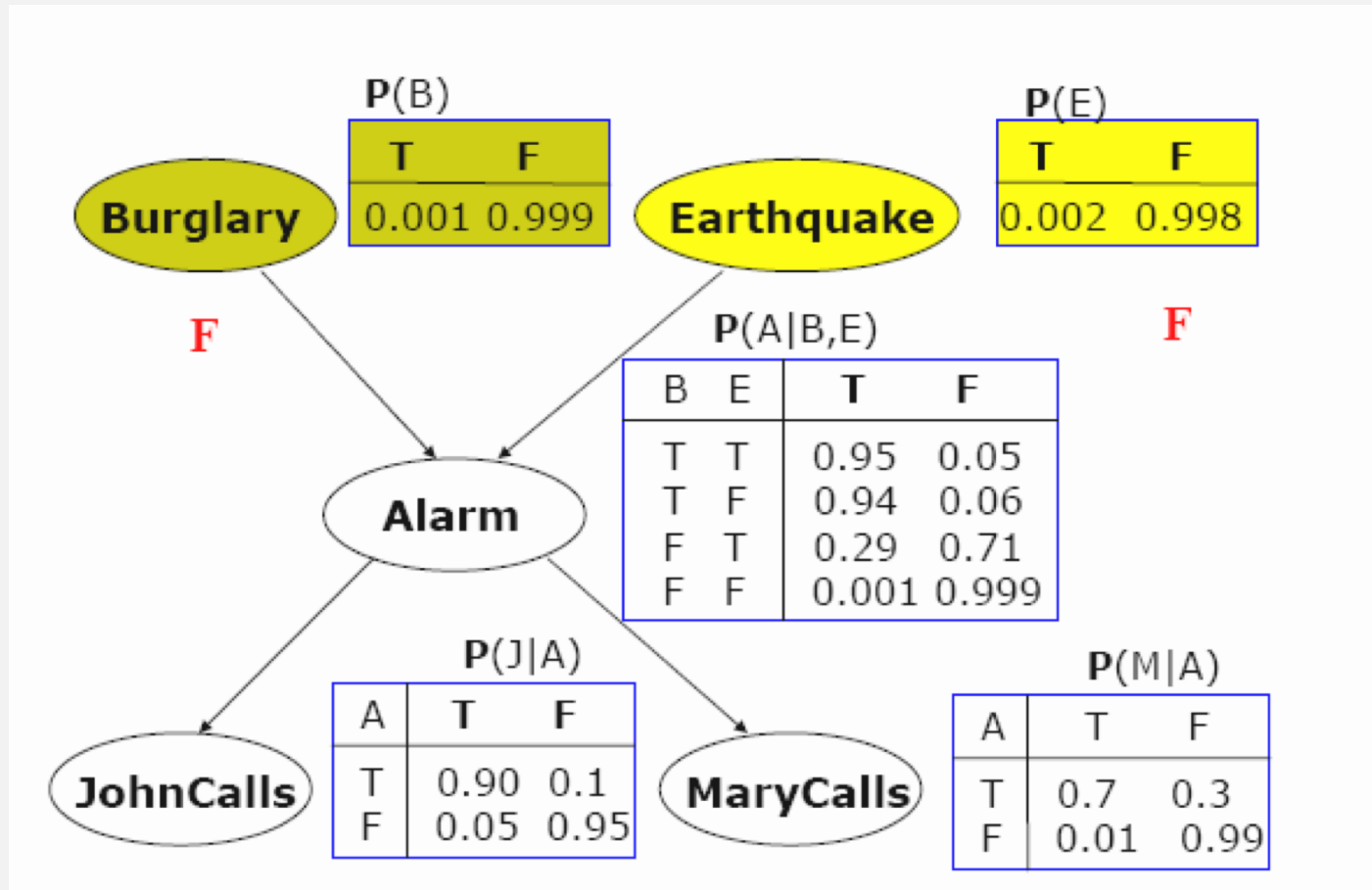
$$P(B | J = T, M = F)$$



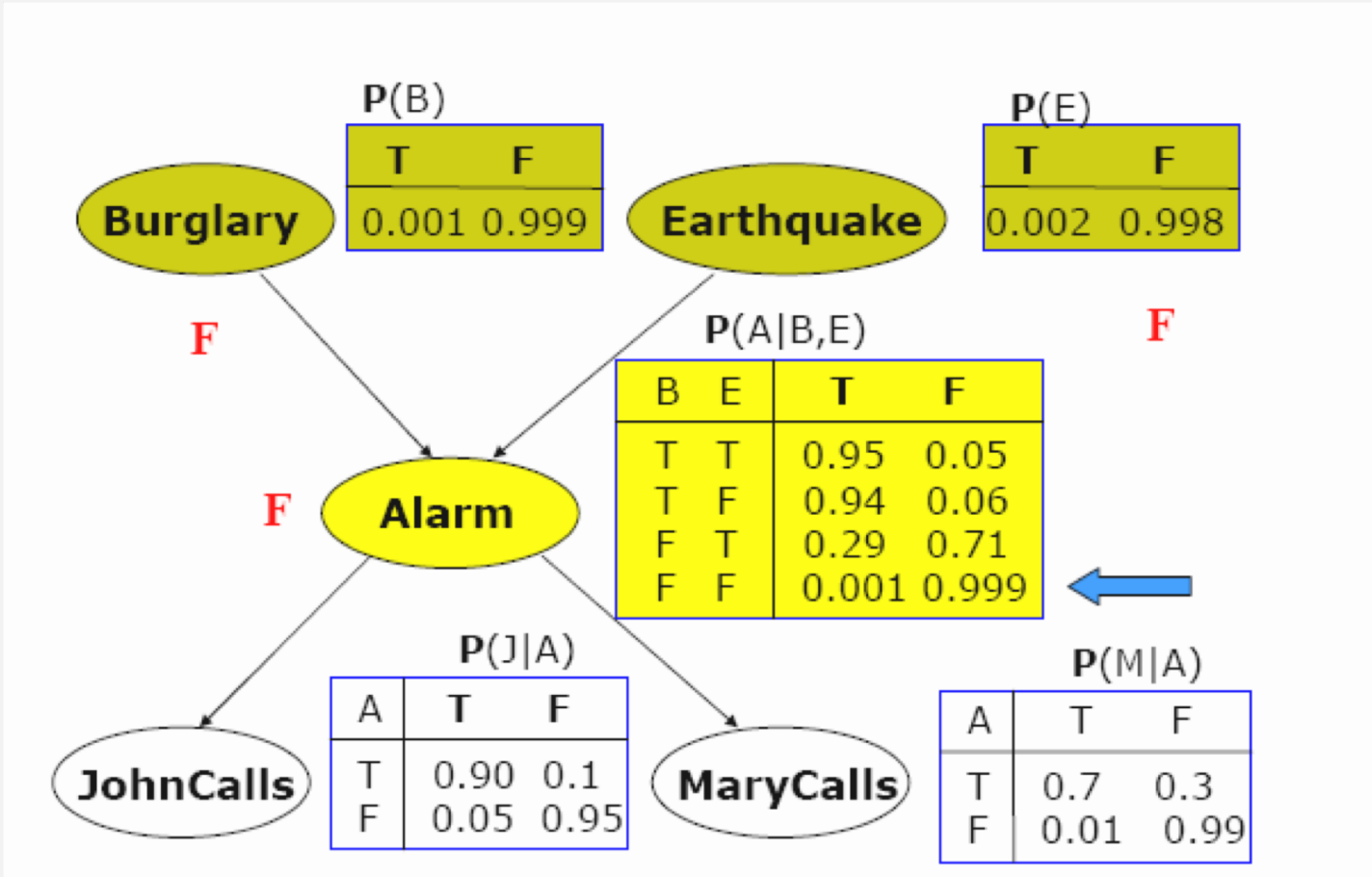
BN Sampling Example



BN Sampling Example

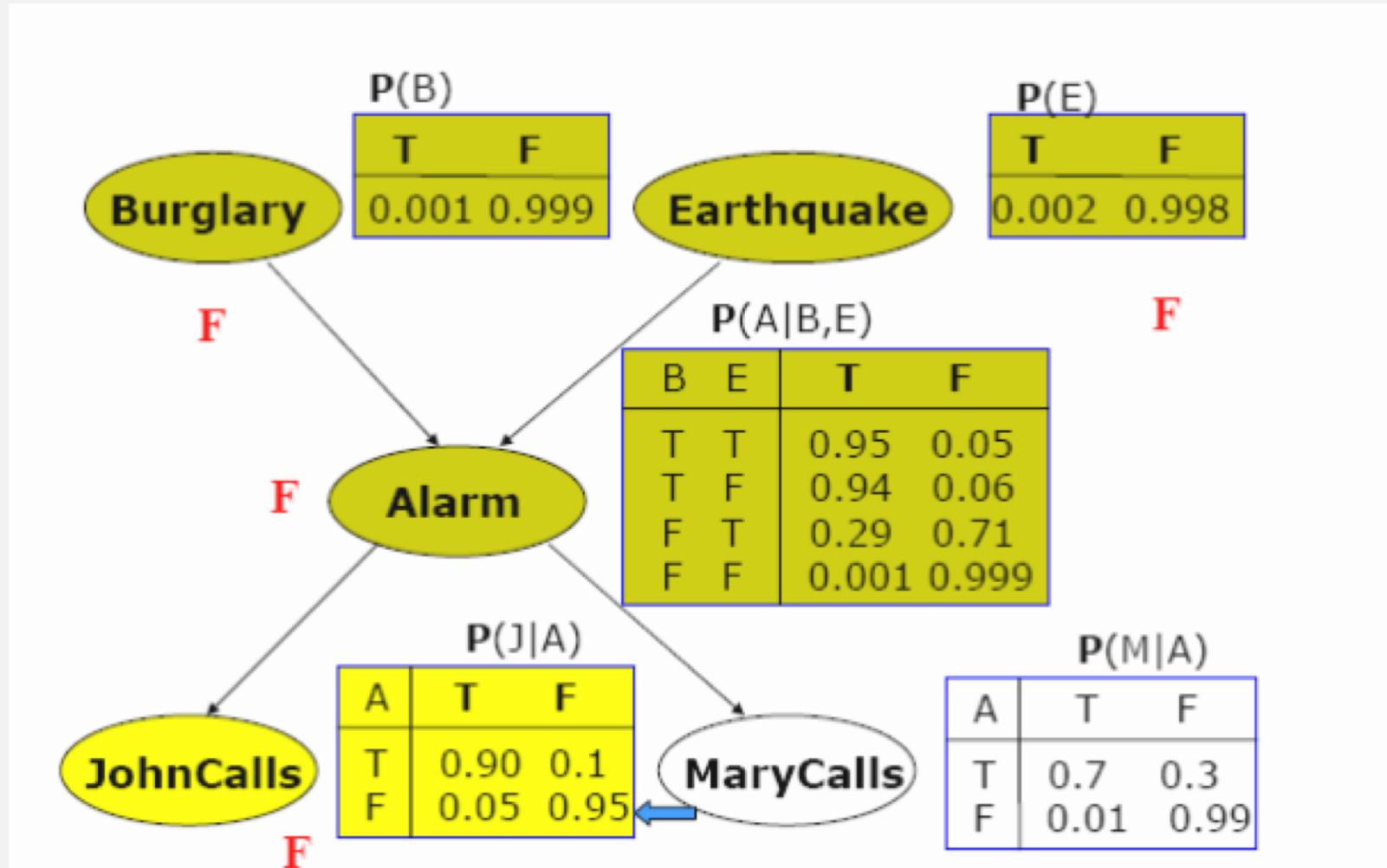


BN Sampling Example

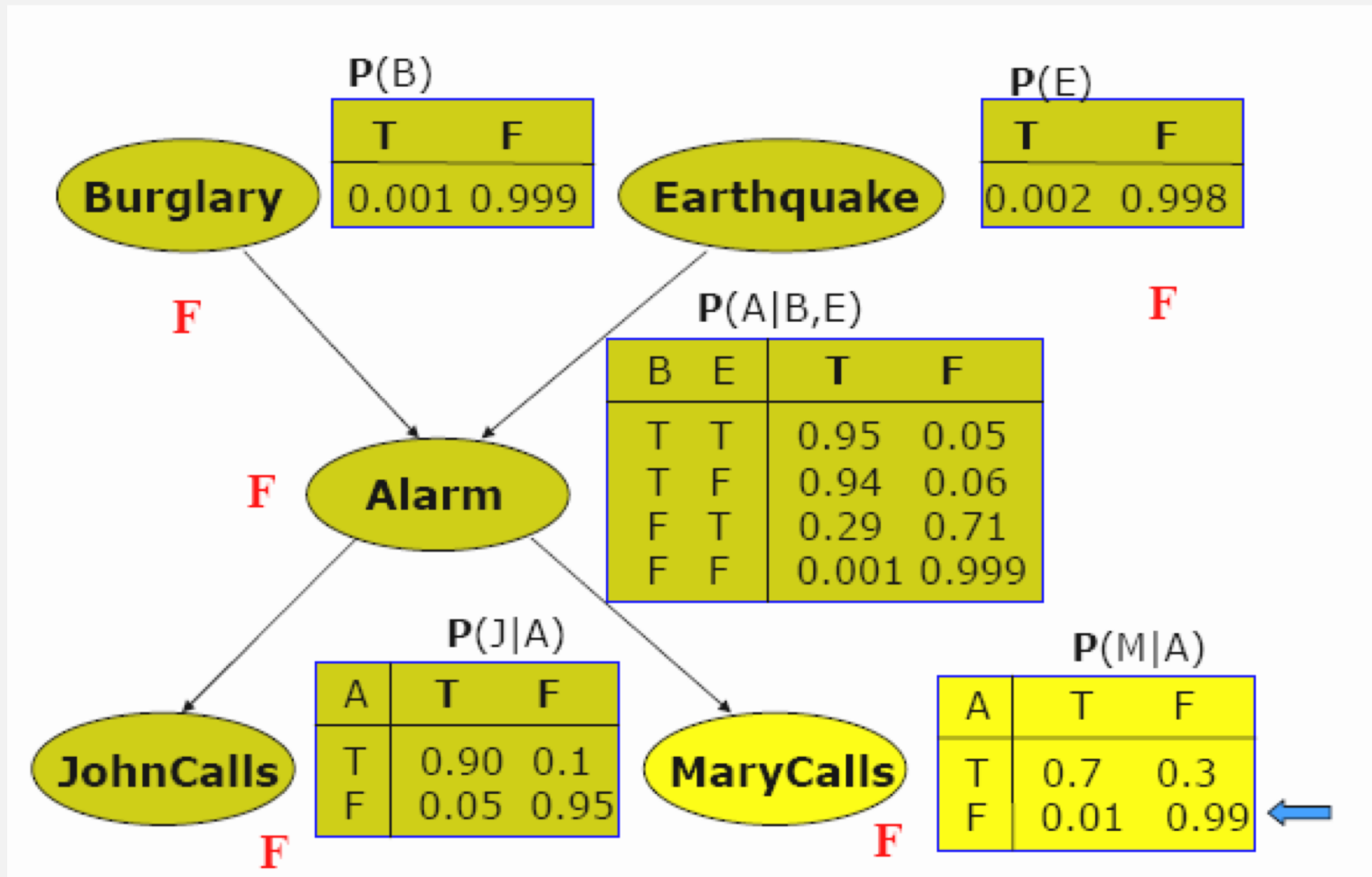




BN Sampling Example



BN Sampling Example





Rejection Sampling

Rejection sampling:

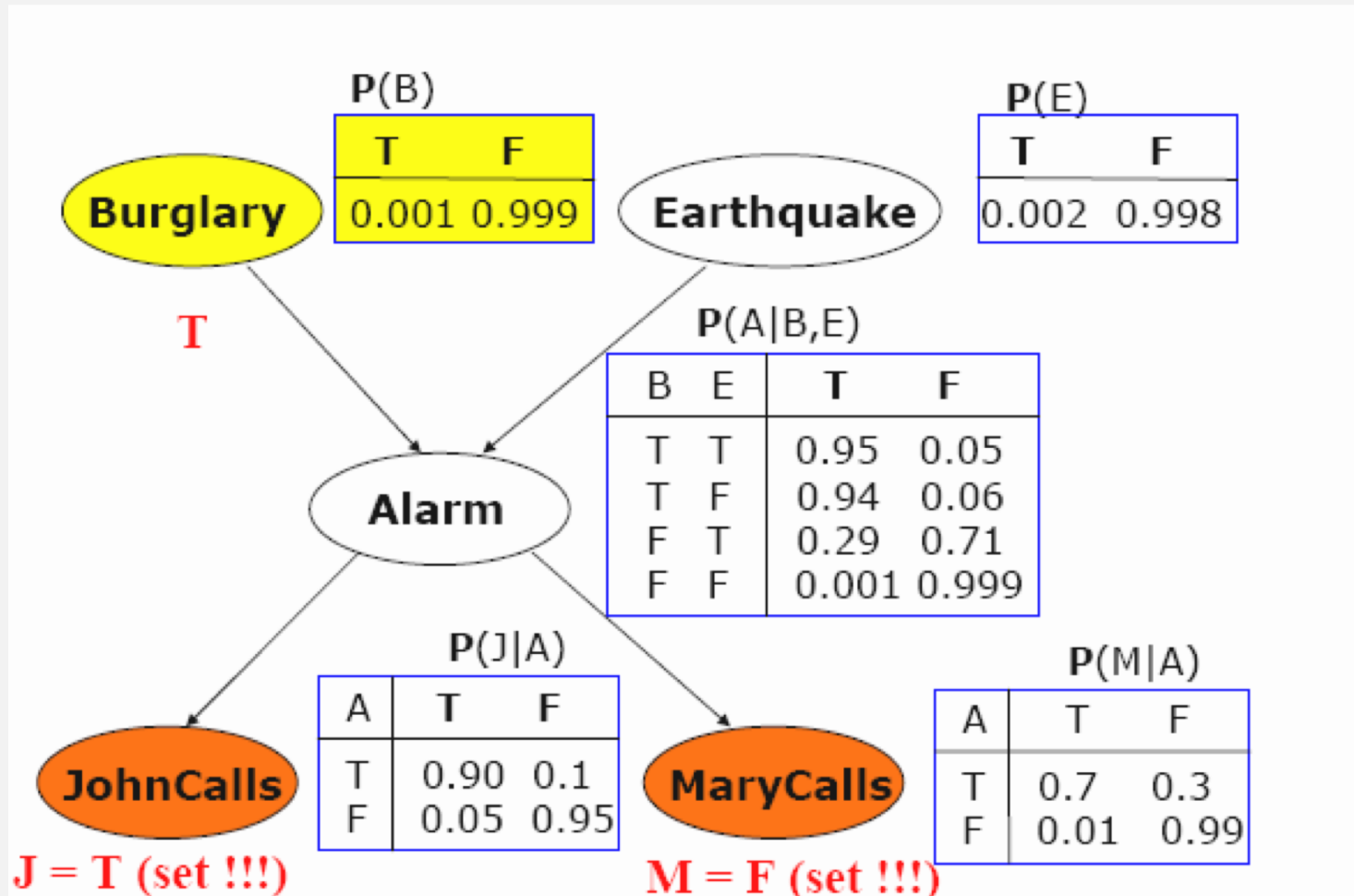
- Generate sample for the full joint by sampling BN
- Use only samples that agree with the condition, the remaining samples are rejected
- Problem: many samples can be rejected



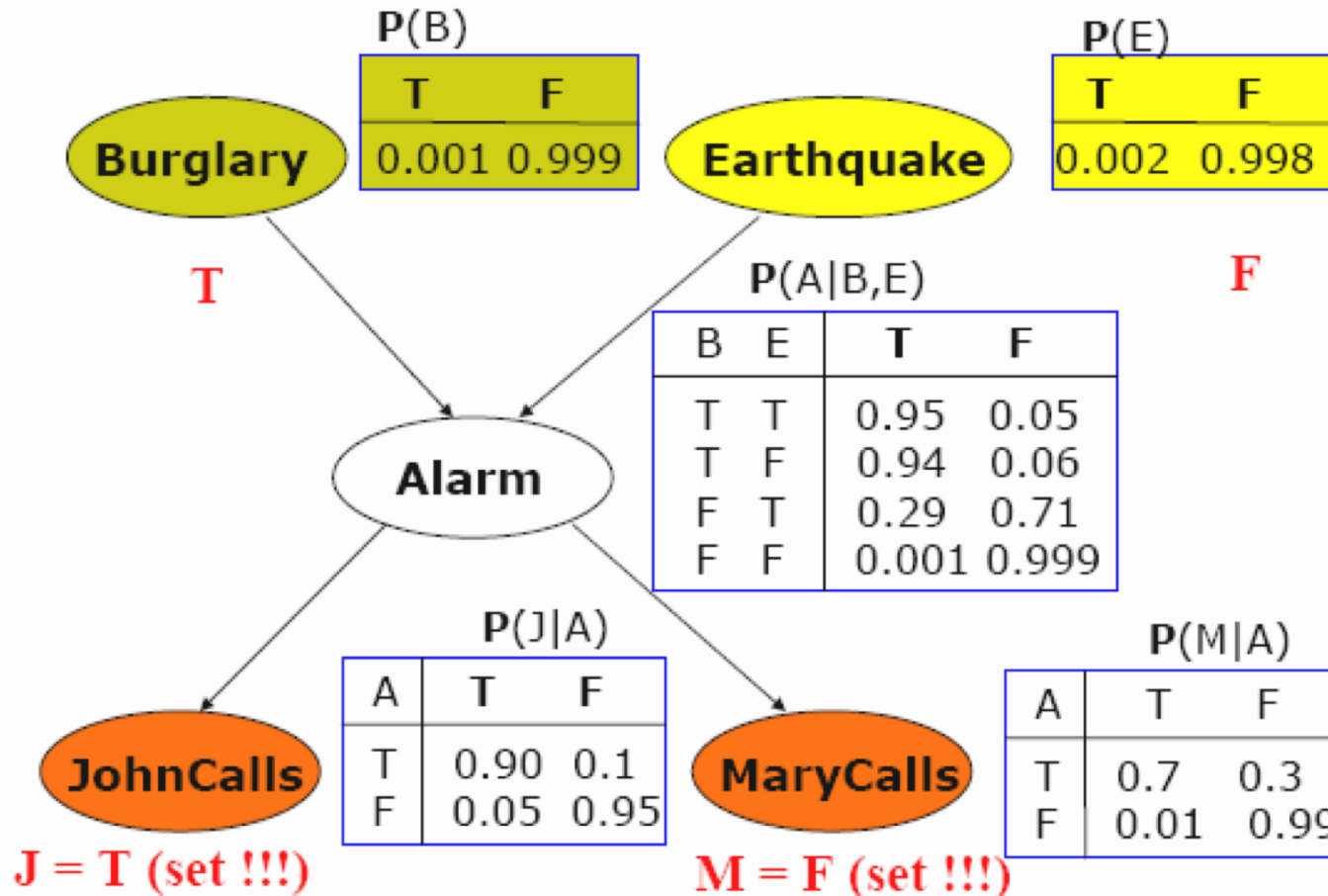
Likelihood weighting

- Avoids inefficiencies of rejection sampling
- Idea: generate only samples consistent with an evidence (or conditioning event)
- If the value is set by evidence, there is no sampling
- Problem: using simple counts is not enough since these may occur with different probabilities
- Likelihood weighting: with every sample keep a weight with which it should count towards the estimate

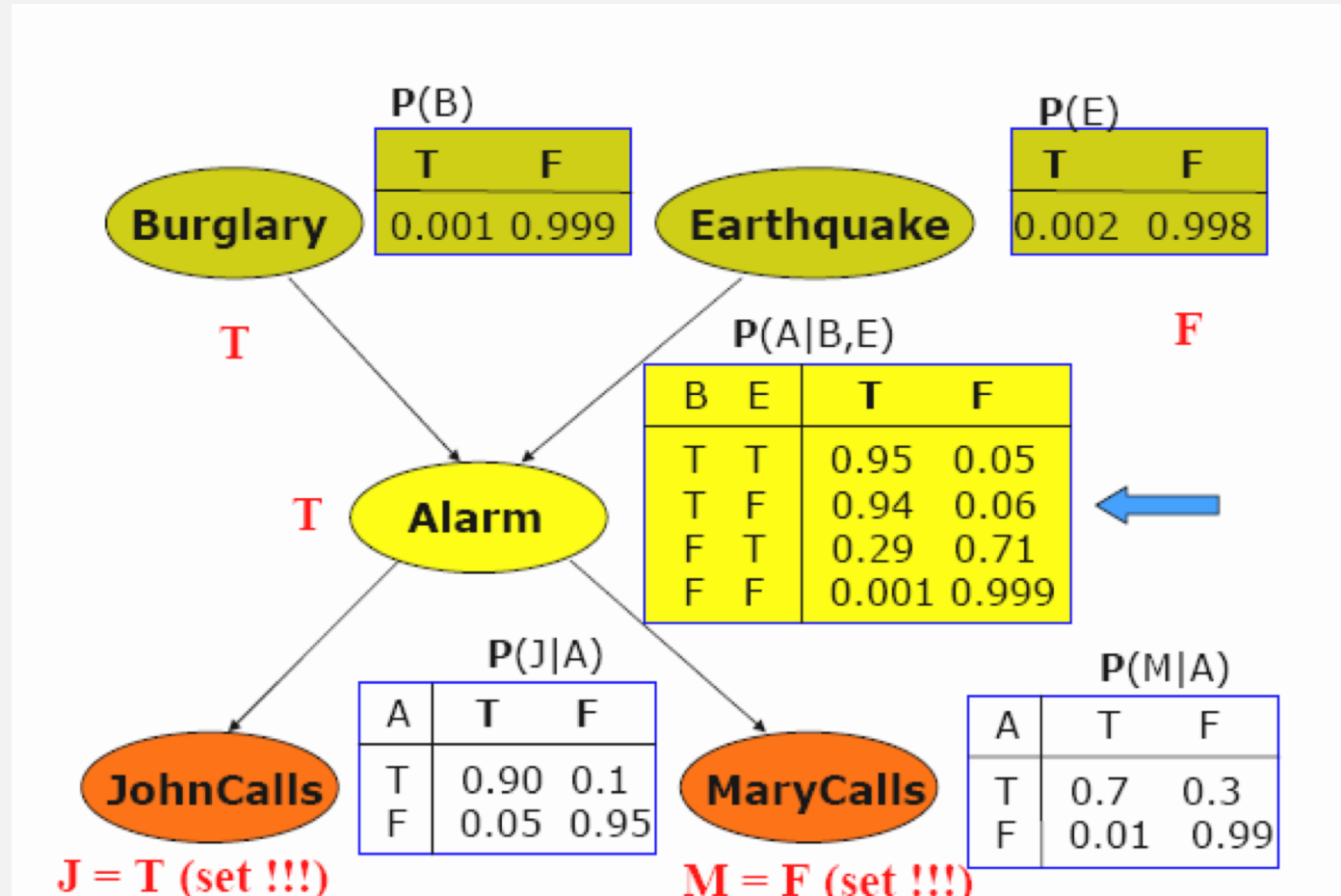
Likelihood weighting Example



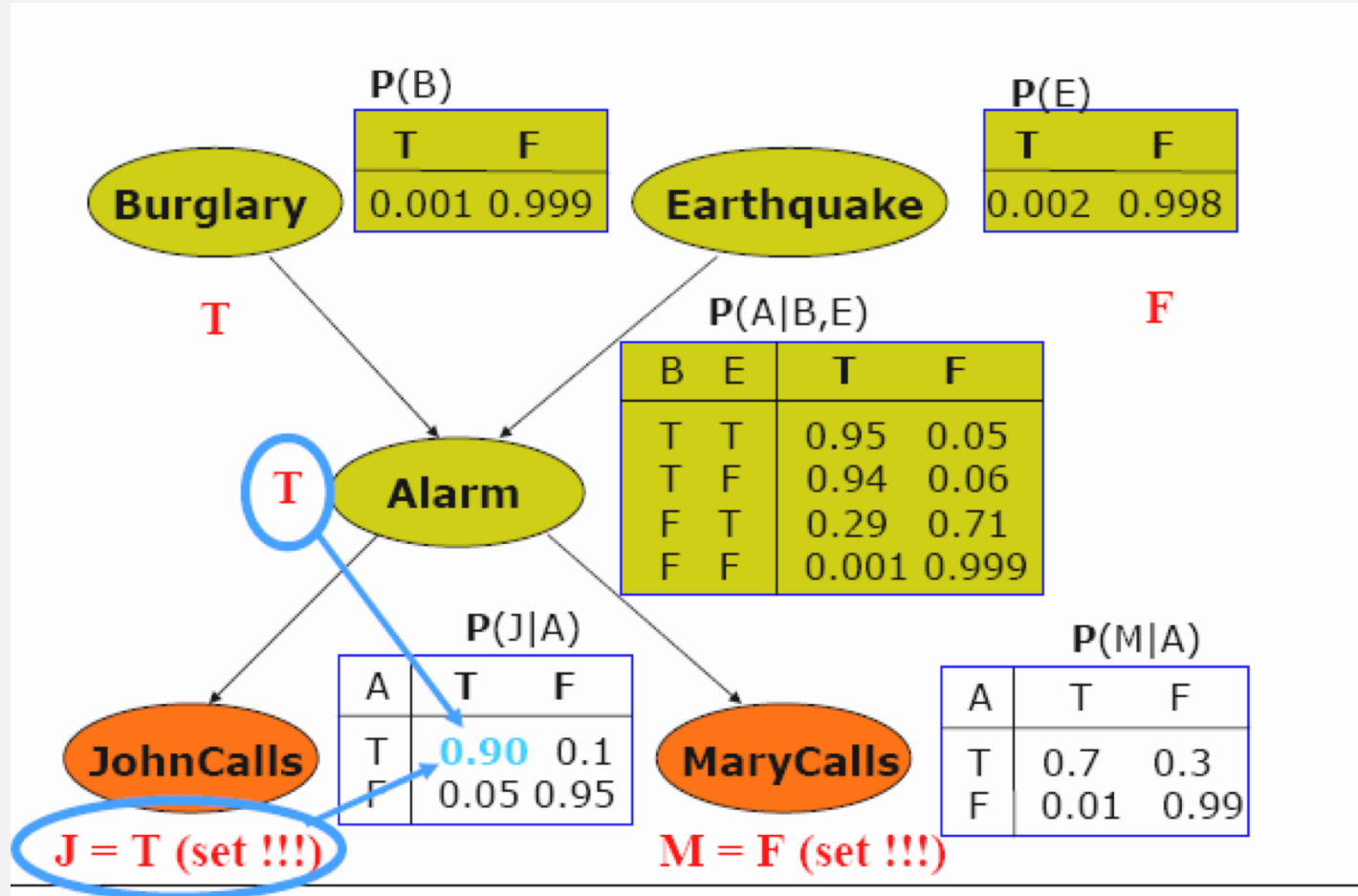
Likelihood weighting Example



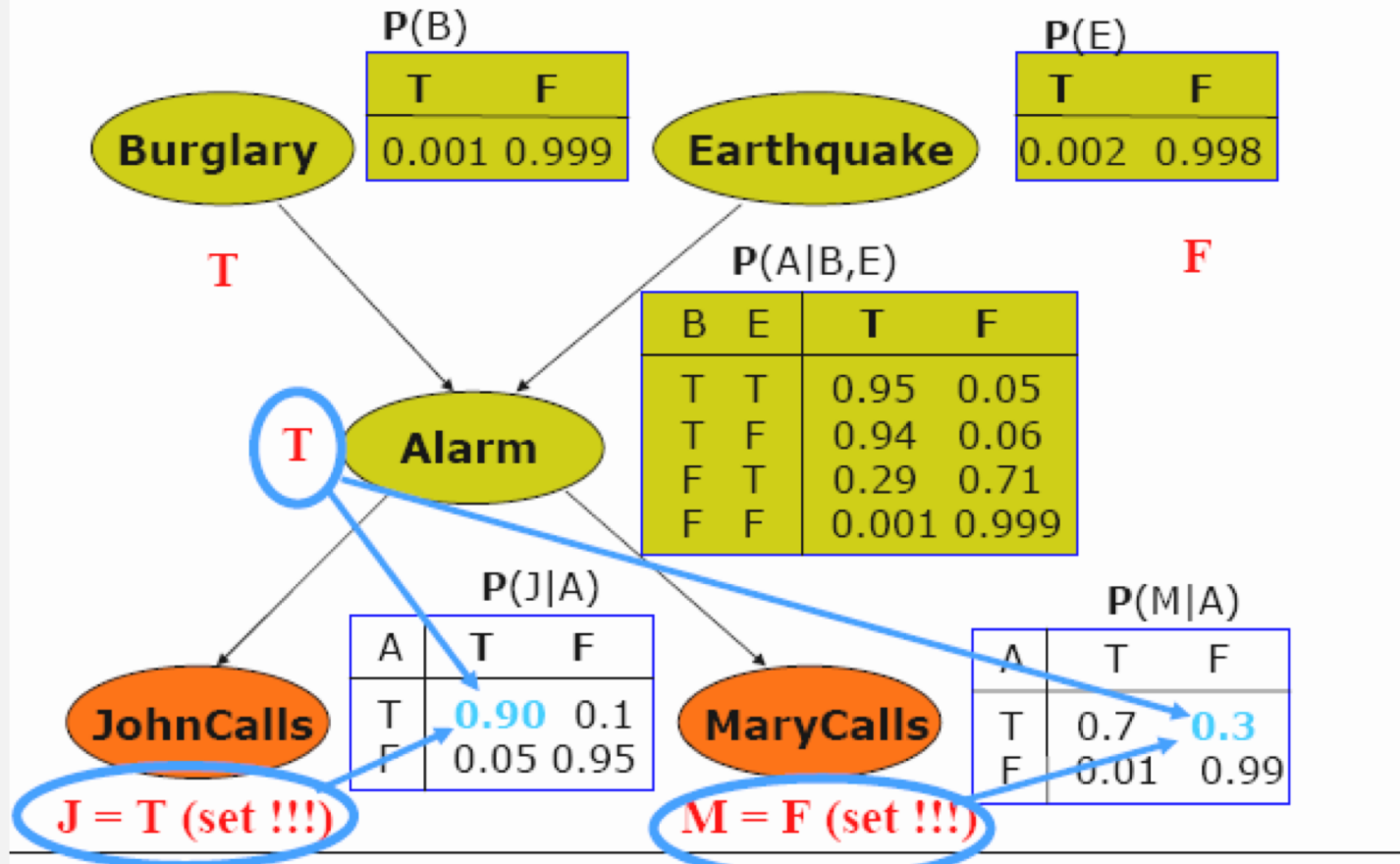
Likelihood weighting Example



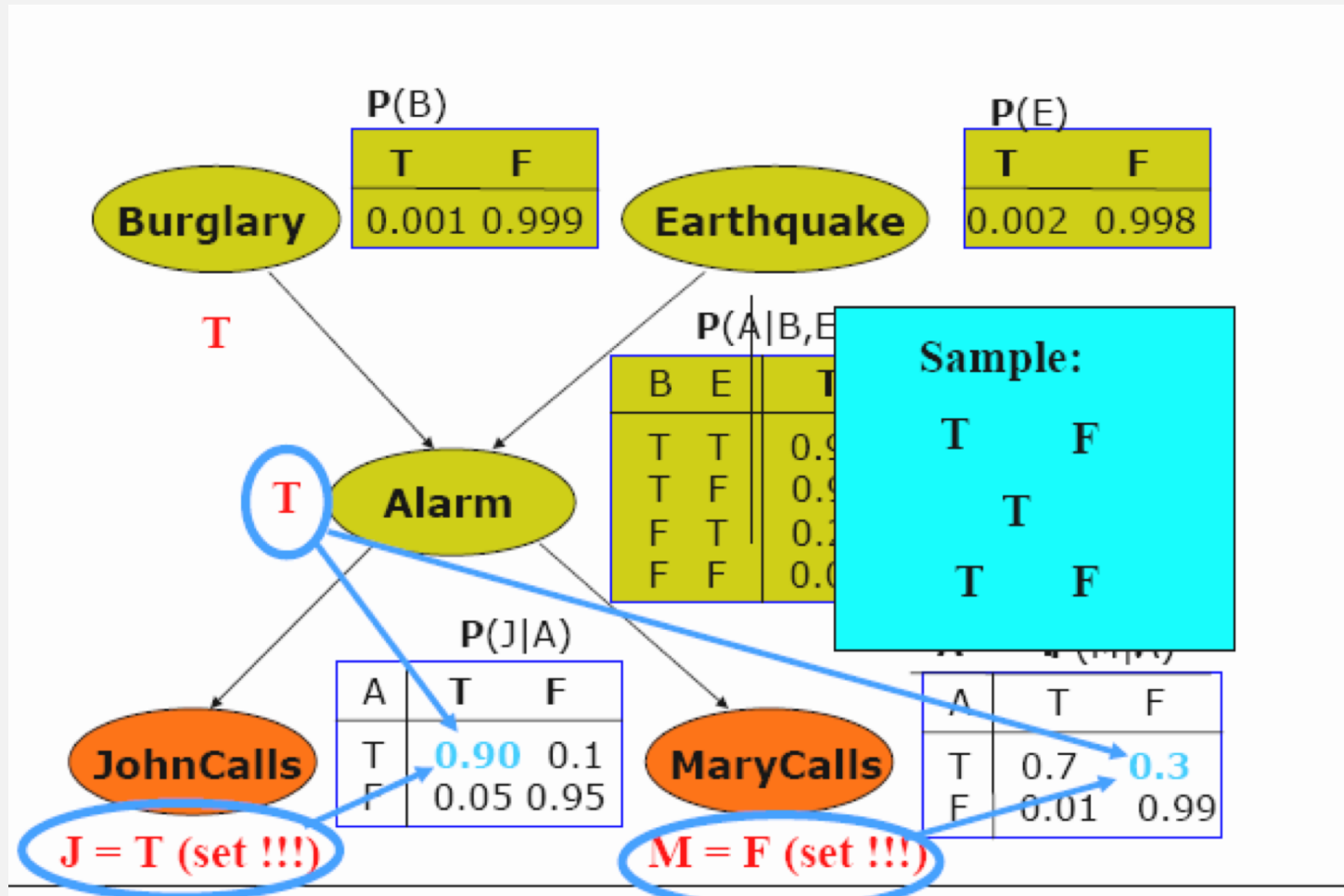
Likelihood weighting Example



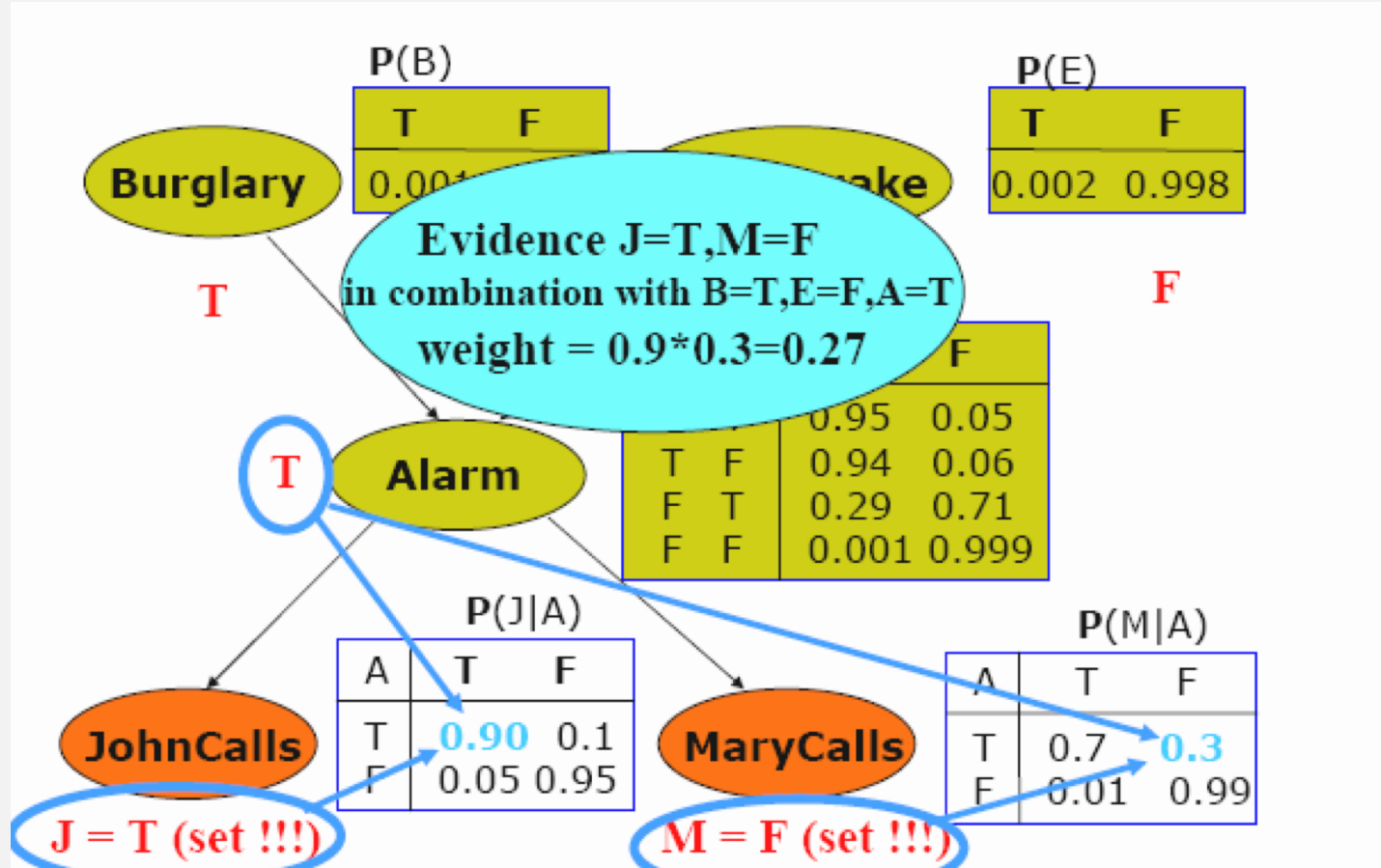
Likelihood weighting Example



Likelihood weighting Example

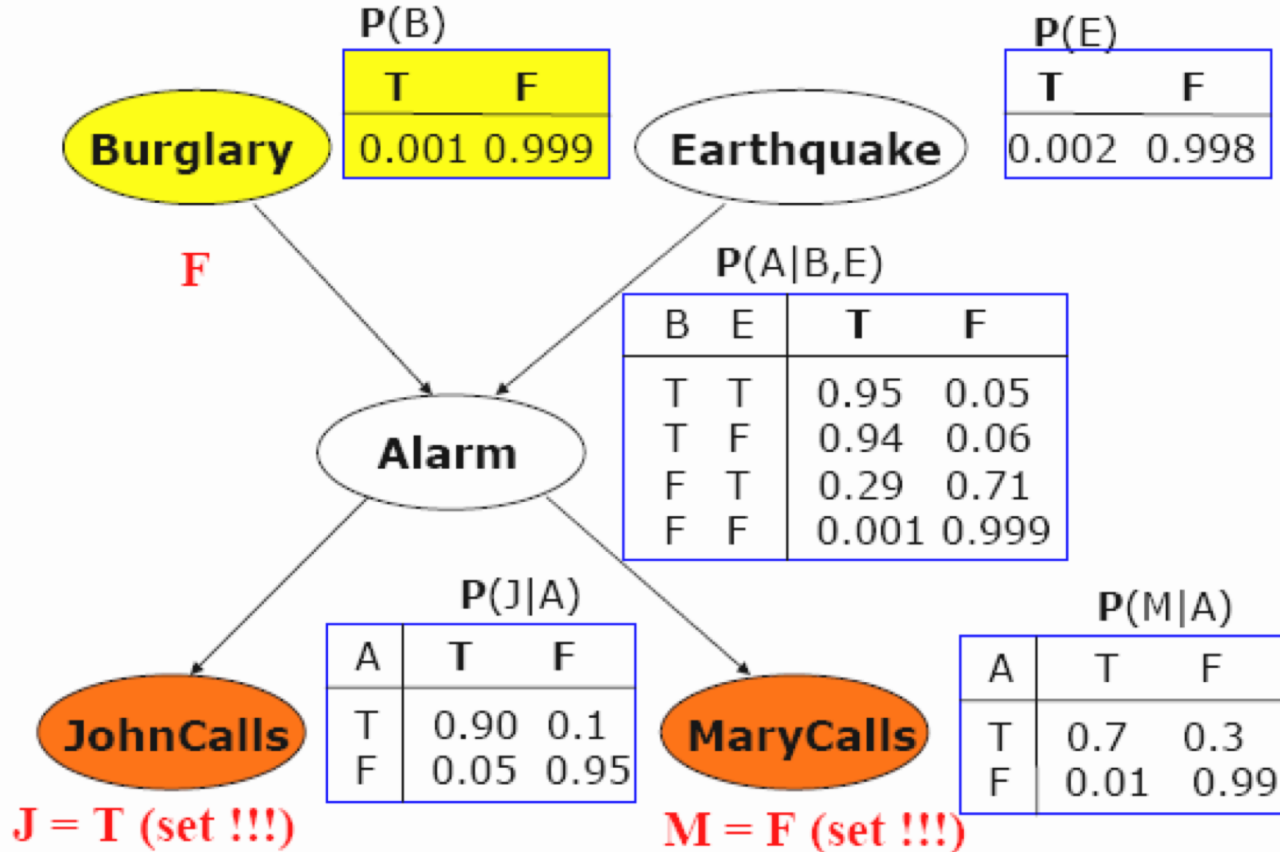


Likelihood weighting Example



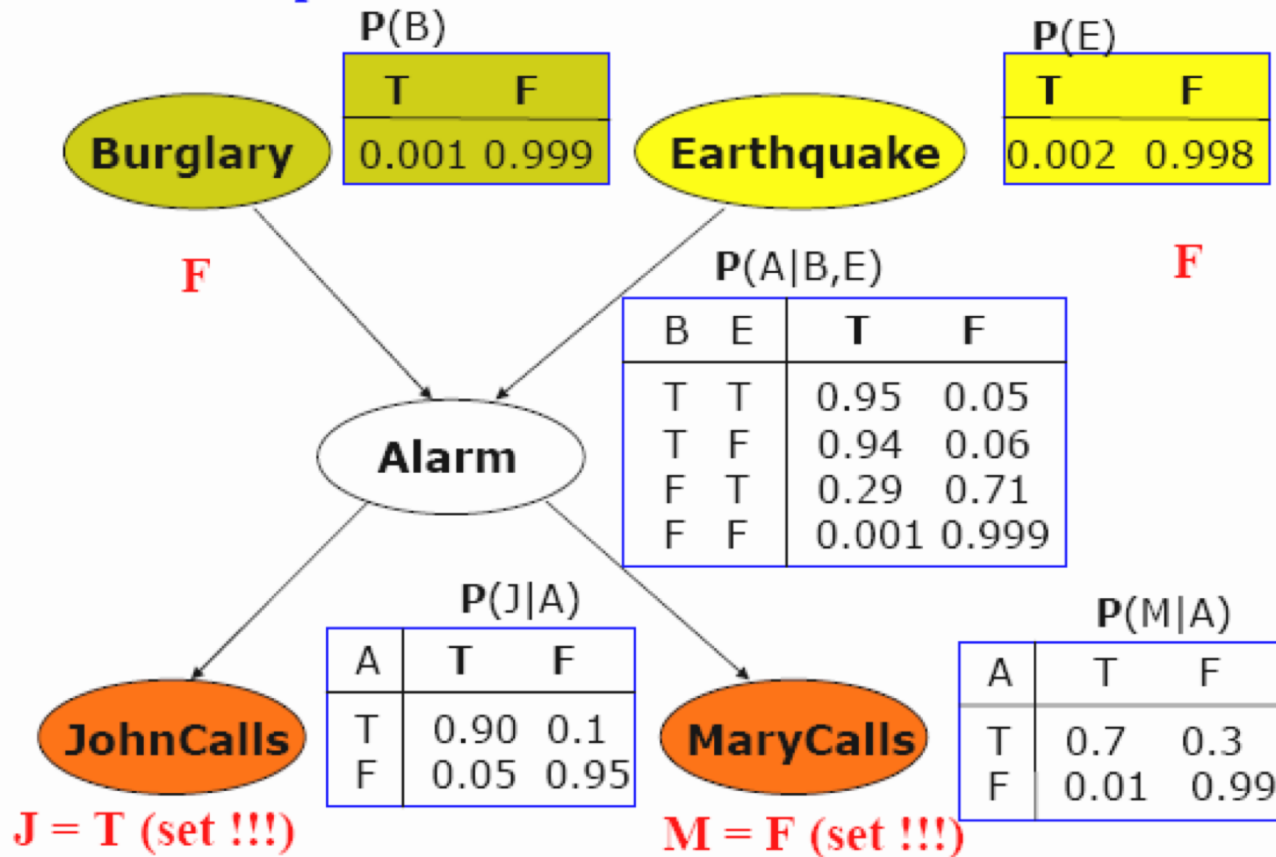
Likelihood weighting Example

Second sample

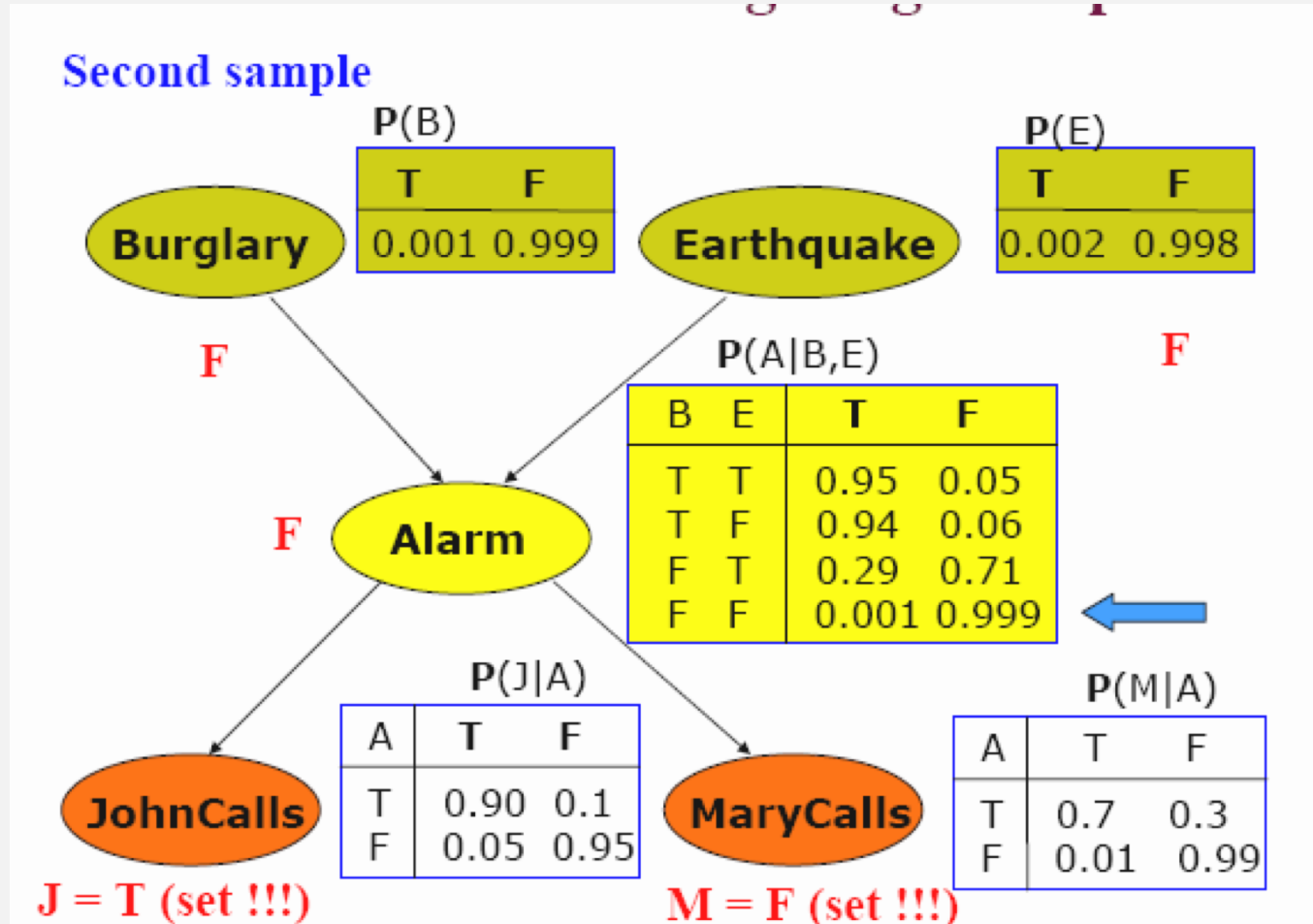


Likelihood weighting Example

Second sample

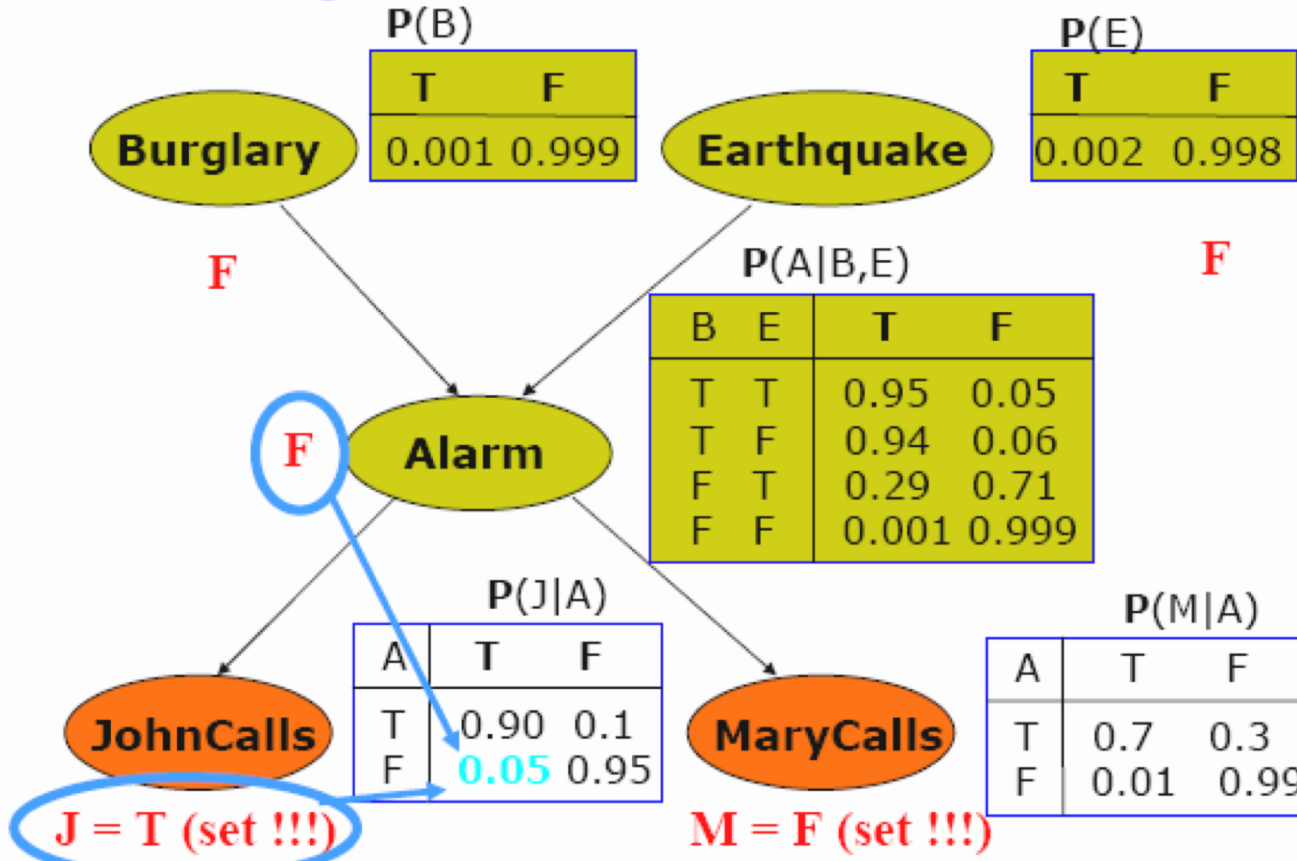


Likelihood weighting Example

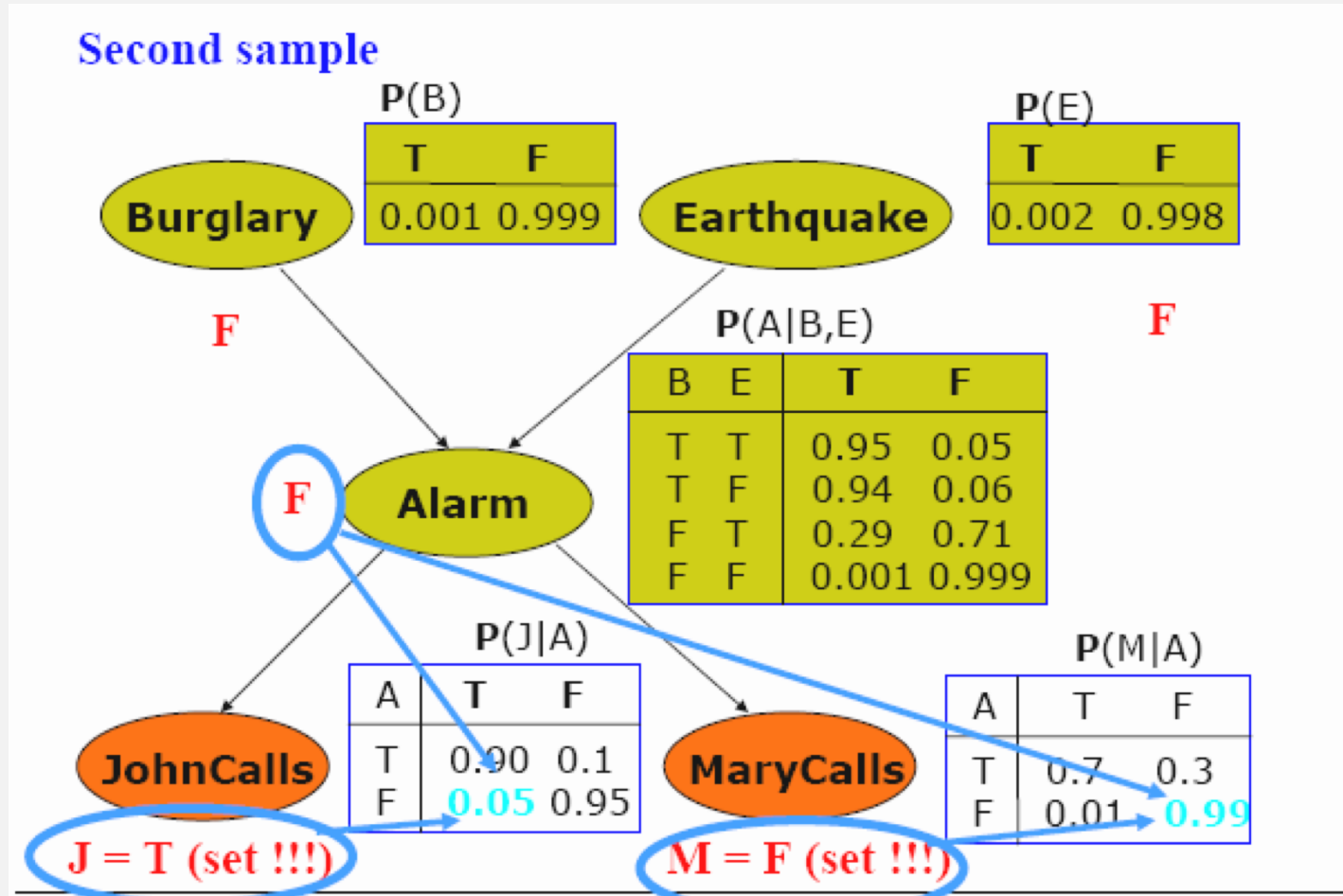


Likelihood weighting Example

Second sample

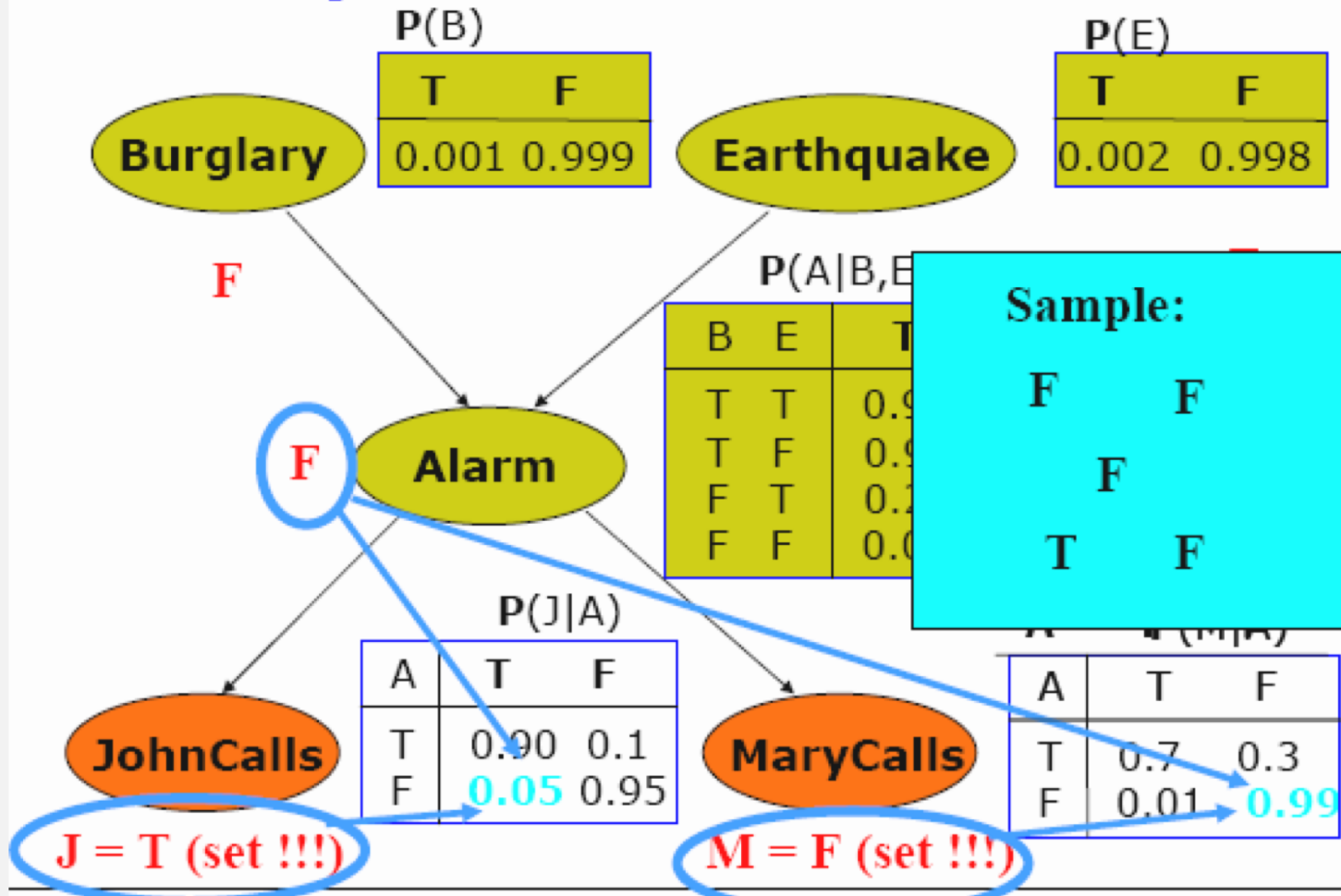


Likelihood weighting Example



Likelihood weighting Example

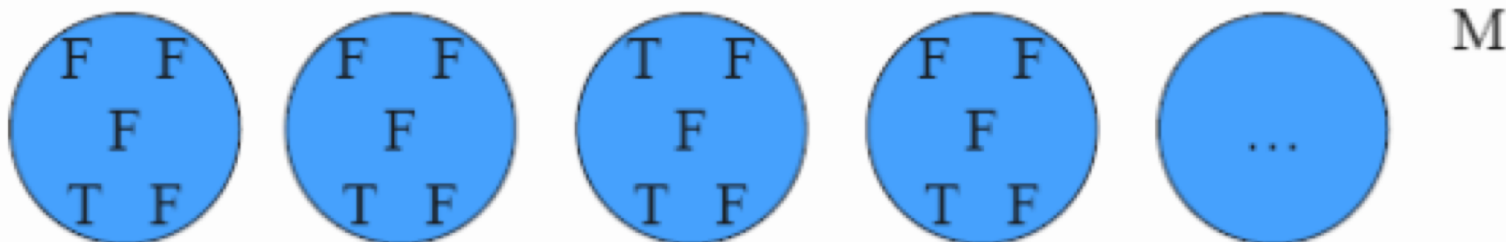
Second sample





Likelihood Sampling

- Assume we have generated the following M samples:



- If we calculate the estimate:

$$P(B=T | J=T, M=F) = \frac{\#sample_with(B=T)}{\#total_sample}$$

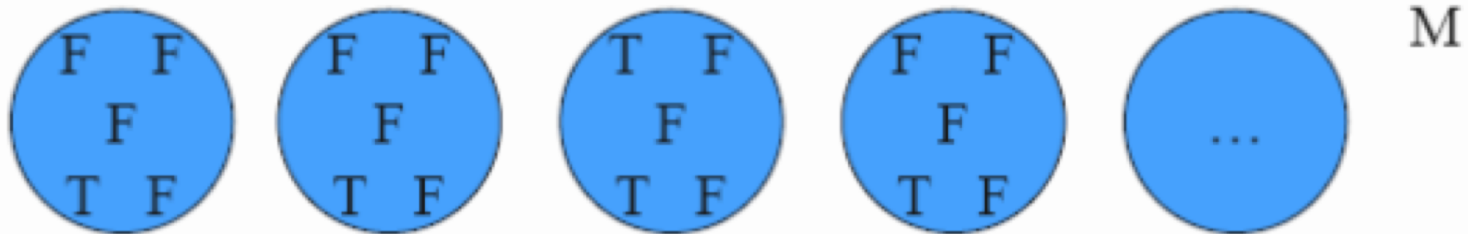
a less likely sample from $P(X)$ may be generated more often.

- For example, sample  is generated more often than in $P(X)$

- So the samples are not consistent with $P(X)$.

Likelihood Sampling

- Assume we have generated the following M samples:



How to make the samples consistent?

Weight each sample by probability with which it agrees with the conditioning evidence $P(e)$.





Likelihood Weighting

- How to compute weights for the sample?
- Assume the query $P(B = T \mid J = T, M = F)$
- Likelihood weighting:
 - **With every sample keep a weight with which it should count towards the estimate**

$$\tilde{P}(B = T \mid J = T, M = F) = \frac{\sum_{i=1}^M 1\{B^{(i)} = T\} w^{(i)}}{\sum_{i=1}^M w^{(i)}}$$

$$\tilde{P}(B = T \mid J = T, M = F) = \frac{\sum_{\text{samples with } B=T \text{ and } J=T, M=F} w_{B=T}}{\sum_{\text{samples with any value of } B \text{ and } J=T, M=F} w_{B=x}}$$



First order probability models

- Can we combine probability with the expressive power of first order logic (FOL) representation?
- Problem: The set of possible worlds represented by an FOL sentence can be infinite
- Relational probability models (RPM) 'solve' this problem by replacing standard FOL semantics by database semantics
 - Unique names assumption (e.g., each customer has a unique ID)
 - Domain closure assumption (there are no more objects beyond the ones that have been named)

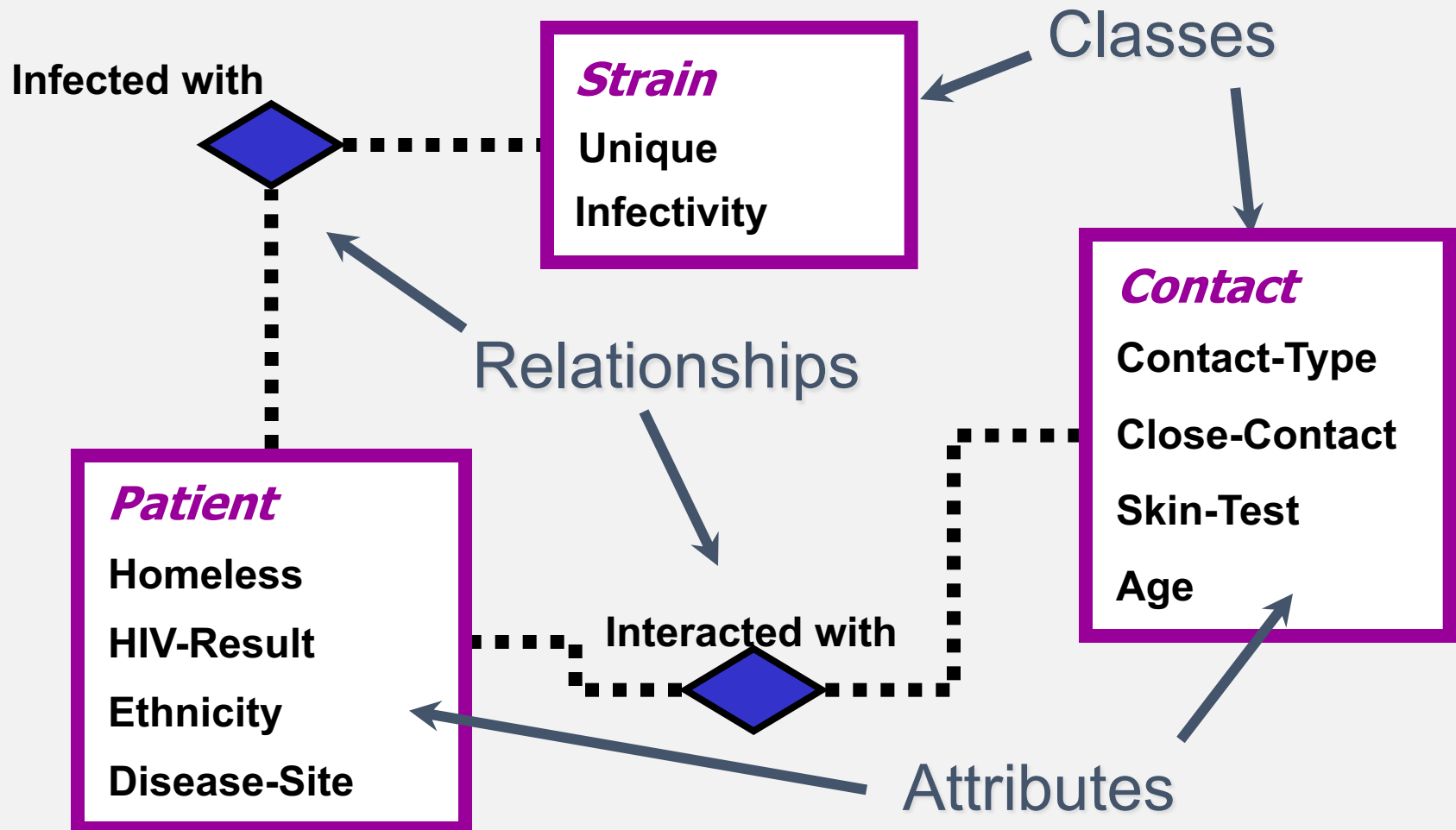
Koller, Pfeffer, Getoor et al. 1999-2007



Probabilistic Relational Models

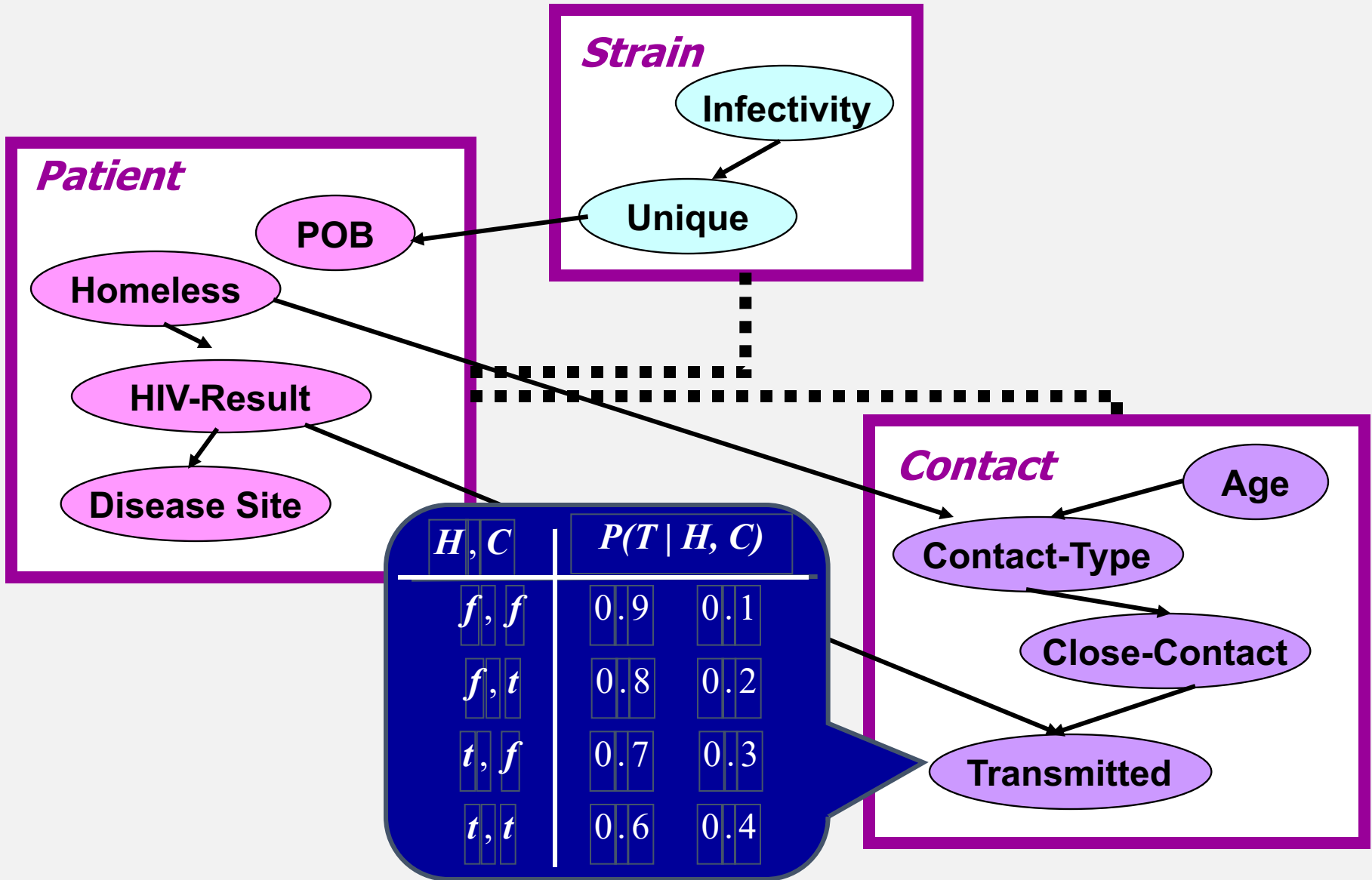
- Combine advantages of relational logic & Bayesian networks:
 - natural domain modeling: objects, properties, relations;
 - generalization over a variety of situations;
 - compact, natural probability models.
- Integrate uncertainty with relational model:
 - properties of entities can depend on properties of related entities;
 - uncertainty over relational structure of domain.

Relational Schema

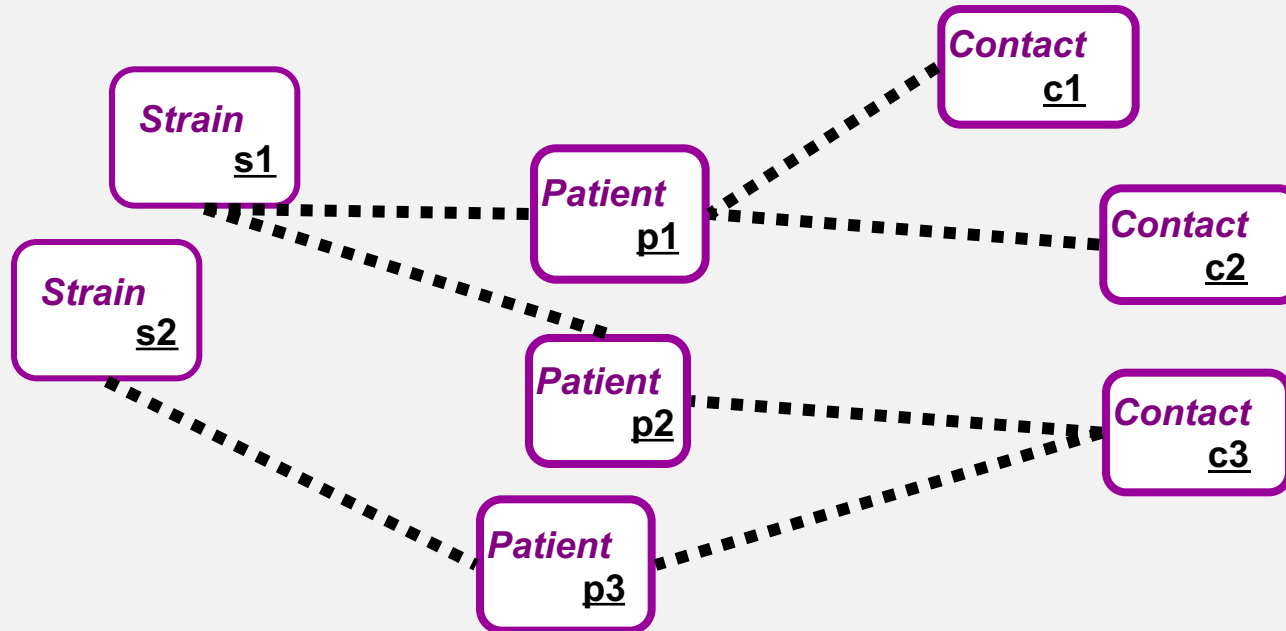


- Describes the types of objects and relations in the database

Probabilistic Relational Model



Relational Skeleton



Fixed relational skeleton σ

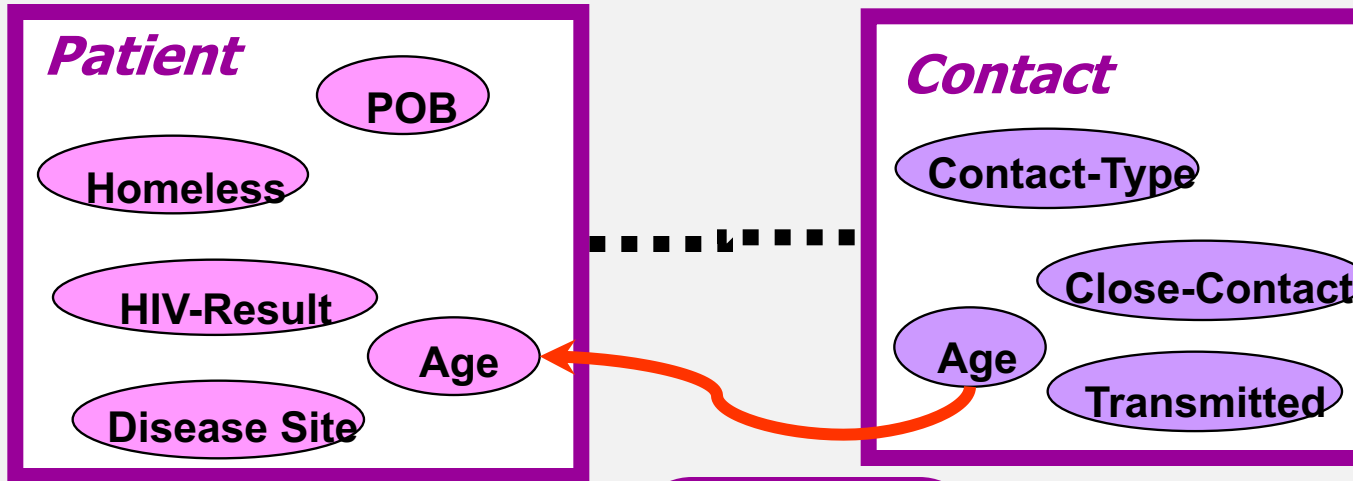
- set of objects in each class
- relations between them

Uncertainty over assignment of values to attributes

PRM defines distribution over instantiations of attributes



PRM: Aggregate Dependencies

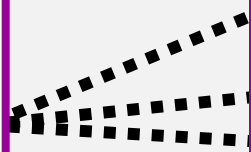


Patient
Jane Doe
POB US
Homeless no
HIV-Result negative
Age ???
Disease Site pulmonary

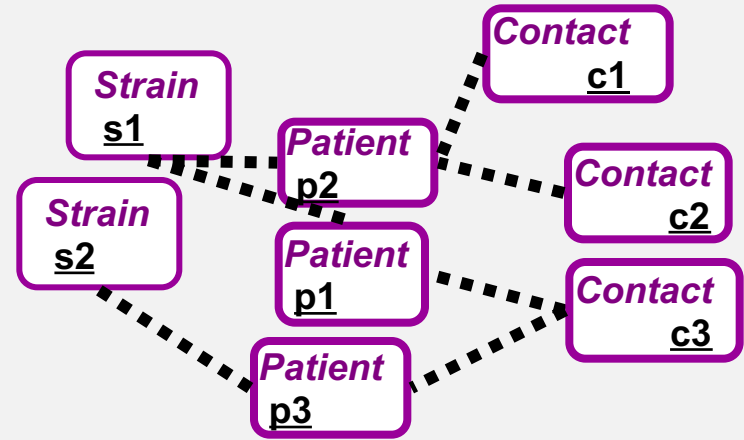
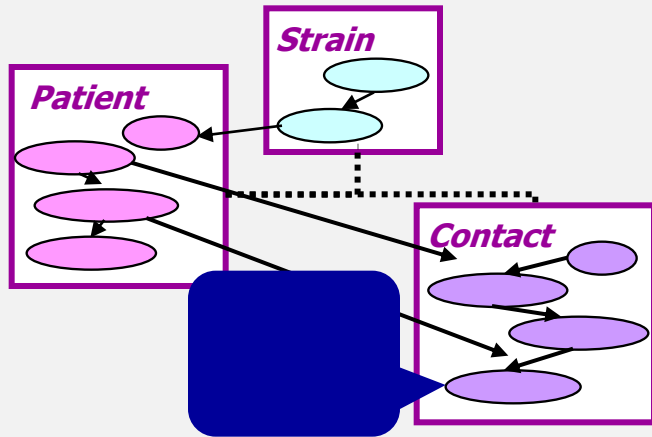
Contact
#5077
Contact-Type coworker
Close-Contact no
Age middle-aged
Transmitted false

Contact
#5076
Contact-Type spouse
Close-Contact yes
Age middle-aged
Transmitted true

Contact
#5075
Contact-Type friend
Close-Contact no
Age middle-aged
Transmitted false



PRM with AU Semantics



PRM + relational skeleton σ =

probability distribution over completions I :

$$P(I \mid \sigma, S, \Theta) = \prod_{x \in \sigma} \prod_{x.A} P(x.A \mid \text{parents}_{S, \sigma}(x.A))$$

↑ Objects ↑ Attributes



Open universe probability models

- Unique names assumption and domain closure assumption do not hold in the presence of uncertainty about existence and identity of objects
- Open universe probability models (OUPMs) extend Bayes networks and RPMs by adding
 - generative steps that add objects to the possible world under construction
 - where the number and type of objects added may depend on the objects that are already present

Milch et al., 2007

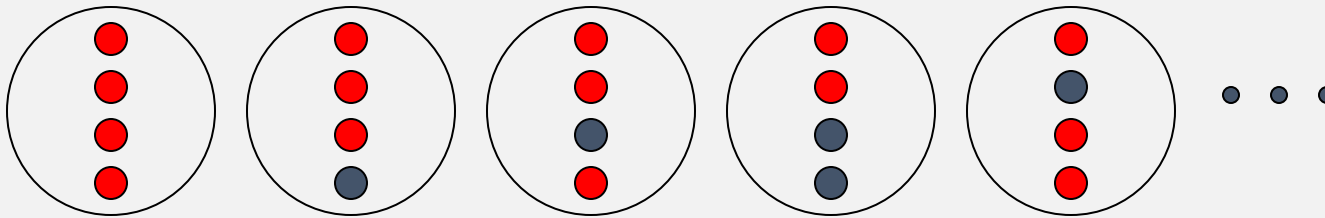


Herbrand vs full first-order semantics

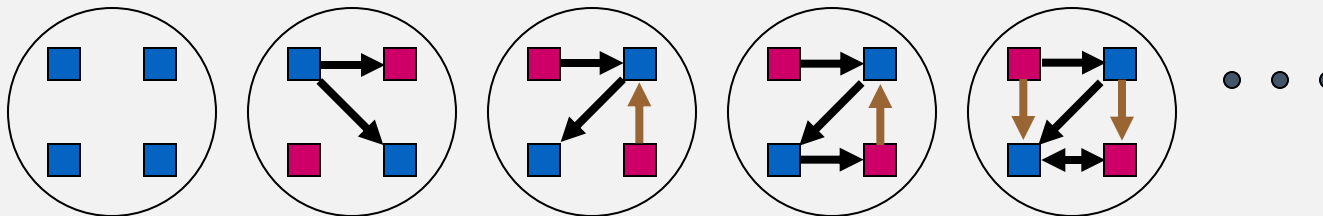
- Given: **Father(Bill,William)** and **Father(Bill,Junior)**
- How many children does Bill have?
 - Database (Herbrand) semantics: 2
 - First-order open world logical semantics:
 - Between 2 and ∞ (under the unique names assumption)
 - Between 1 and ∞ (in the absence of the unique names assumption)

Possible worlds

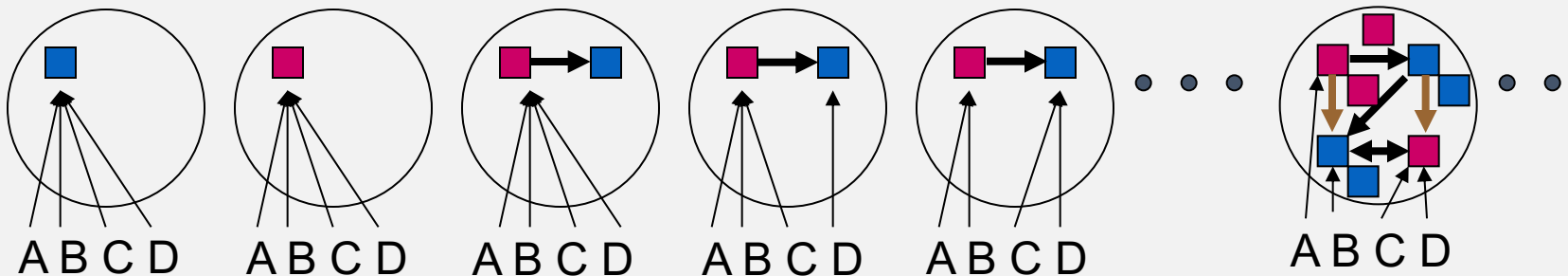
- Propositional (Boolean logic, Bayes nets)



- First-order closed-universe (DB, RPM)



- First-order open-universe: uncertainty about existence of objects and the relations





Open-universe models in BLOG

- Construct worlds using two kinds of steps, proceeding in topological order:
 - **Dependency statements**: Set the value of a **function or relation on a tuple of (quantified) arguments**, conditioned on parent values



Open-universe models in BLOG

- Construct worlds using two kinds of steps, proceeding in topological order:
 - **Dependency statements**: Set the value of a **function or relation on a tuple of (quantified) arguments**, conditioned on parent values
 - **Number statements**: **Add some objects to the world**, conditioned on what objects and relations exist so far



Technical basics

Theorem: Every well-formed* BLOG model specifies a unique proper probability distribution over open-universe possible worlds; equivalent to an infinite contingent Bayes net

Theorem: BLOG inference algorithms (rejection sampling, importance sampling, MCMC) converge to correct posteriors for any well-formed* model, for any first-order query



Example: cyber-security sibyl defense

```
#Person ~ LogNormal[6.9, 2.3]();  
Honest(x) ~ Boolean[0.9]();  
#Login(Owner = x) ~  
  if Honest(x) then 1 else LogNormal[4.6,2.3]();  
Transaction(x,y) ~  
  if Owner(x) = Owner(y) then SibylPrior()  
  else TransactionPrior(Honest(Owner(x)),  
                        Honest(Owner(y)));  
Recommends(x,y) ~  
  if Transaction(x,y) then  
    if Owner(x) = Owner(y) then Boolean[0.99]()  
    else RecPrior(Honest(Owner(x)),  
                 Honest(Owner(y)));
```

Evidence: lots of transactions and recommendations

Query: Honest(x)





Probabilistic Programming Languages

- Logic based
 - PRISM, Problog – logic programming + probability distributions over facts [Sato and Kameya, 2001; De Raedt, Kimmig, and Toivonen, 2007]
 - BLOG – a language based on open universe probability models [Milch et al., 2007]
- Functional programming based
 - Church, Venture – extend Scheme with probabilistic semantics for specifying recursively defined generative processes [Goodman, Mansinghka, Roy, Bonawitz and Tenenbaum, 2008]
 - IBAL – a stochastic functional programming language [Pfeffer, 2007]
- Object-oriented
 - Figaro – an expressive language with support for directed and undirected probabilistic graphical models, OUPMs, models defined over complex data structures. [Pfeffer, 2009]